

Learning Objectives

1 Define a multilevel measurement model (MLMM)

2 Identify when an MLMM is needed

3 Describe and execute an MLMM in a multilevel modelling framework

4 Describe and execute an MLMM in a structural equation modelling framework

Module Overview

- 5 sections of content
 - Review of multilevel modelling
 - Review of measurement modelling
 - Multilevel measurement modelling (MLMM) overview
 - MLMM in MLM framework
 - MLMM in SEM framework
- Code and data for each module available to download

About the Authors

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- Ph.D. student in Quantitative Psychology at McGill University
- Research on effect sizes in multilevel modelling and measurement in replication studies
- Teaches multilevel modelling
 - Co-author of open-source teaching materials at www.learn-mlms.com



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- Assistant Director for Methods at the Psychological Science Accelerator
- Research on measurement practices and the appropriate and transparent use of latent variable models in psychological research
- Teaches courses on introductory statistics, measurement theory, and multilevel modelling
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Shaw, M. & Flake, J. K. (2023). Multilevel Measurement Models [Digital ITEMS Module 34]. *Educational Measurement: Issues and Practice*, 34(4), 82.

Multilevel Modelling Overview

A series of vertical bars of varying heights and colors (maroon, gold, and light purple) on the left side of the slide.

1

A large maroon circle with a gold outline, containing the number 1 in gold.

1

Multilevel Modelling Overview

Section Learning Objectives

Understand when and why to use multilevel models

Recognize when the model is cross-sectional or repeated measures

Write and understand multilevel modelling equations

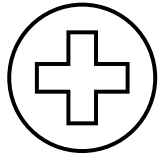
Interpret output for fixed and random effects

Clustered Data Structures

- Common in educational and psychological research



Students within classes



Patients within clinicians



Trials within people (repeated measures)

Why Care About Clustering?



- At least two levels of variance: within and between clusters
 - Level 1 = students
 - Level 2 = class
- What if we neglect clustering?
 - Inflates sample size, deflates standard errors, increases likelihood of Type 1 Error
 - We miss out on interesting multilevel questions!

Example Equations



Outcome: student math achievement (math)

Level 1 predictor: hours spent studying (hours)

Level 2 predictor: teacher's years of experience (texp)

- Random intercept and random slope

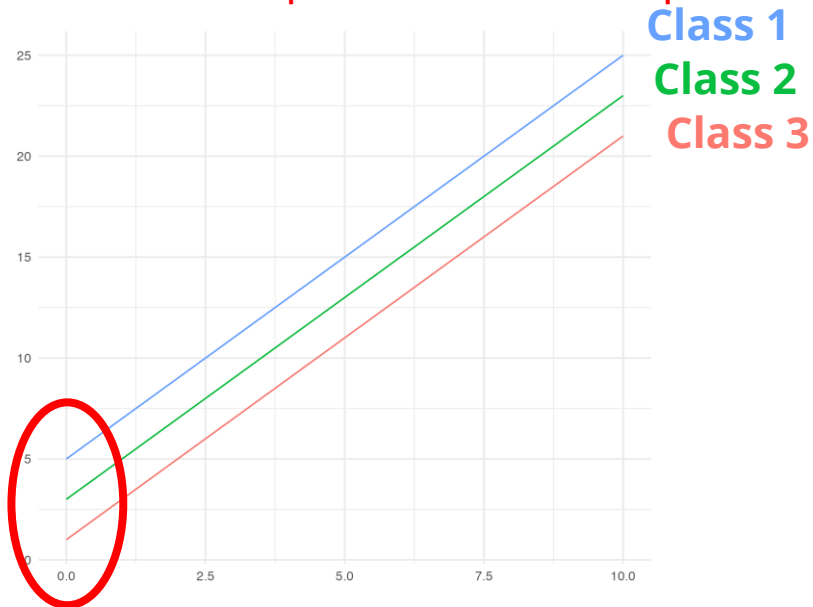
Random Intercept and Random Slope

Outcome: student math achievement (math)

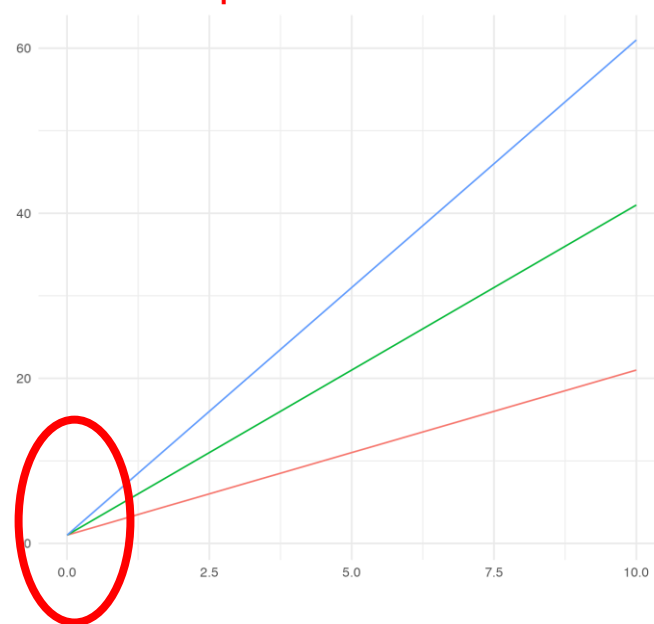
Level 1 predictor: hours spent studying (hours)

Level 2 predictor: teacher's years of experience (texp)

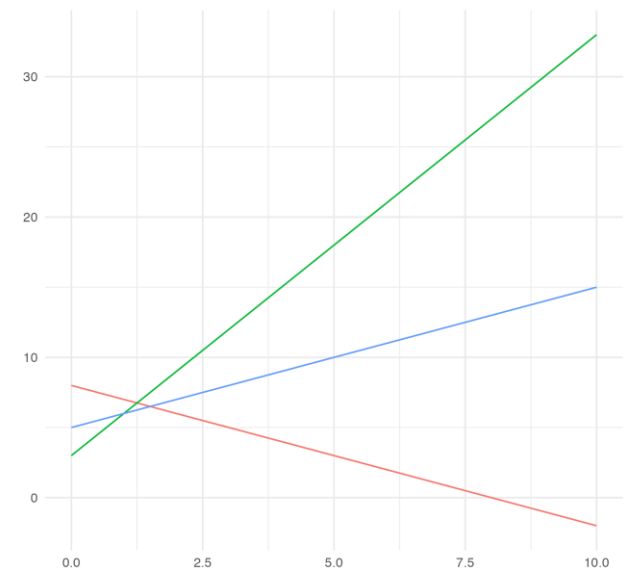
Random intercept, consistent slope



Random slope, consistent intercept



Random slope and intercept



Example Equations

Outcome: student math achievement (math)

Level 1 predictor: hours spent studying (hours)

Level 2 predictor: teacher's years of experience (texp)

- Random intercept and random slope
- Level 1 equation: $math_{ij} = \beta_{0j} + \beta_{1j}hours_{ij} + \epsilon_{ij}$
- Level 2 equations: $\beta_{0j} = \gamma_{00} + \gamma_{01}texp_j + U_{0j}$
- Total: $math_{ij} = \gamma_{00} + \gamma_{01}texp_j + \gamma_{10}hours_{ij} + U_{0j} + U_{1j}hours_{ij} + \epsilon_{ij}$
- Centering requires extra consideration in MLMs
- Without centering, effects are blends of within and between effects

Enders, C., & Tofighi, D. (2007). Centering Predictor Variables in Cross-Sectional Multilevel Models: A New Look at An Old Issue. *Psychological Methods*, 12, 121-138. doi:10.1037/1082-989X.12.2.121

Generalized Equation


Level 1: $y_{ij} = \beta_{0j} + \sum_{p=1}^P \beta_{pj} X_{pij} + \epsilon_{ij}$

Level 2 intercept: $\beta_{0j} = \gamma_{00} + \sum_{q=1}^Q \gamma_{0q} Z_{qj} + U_{0j}$

Level 2 slopes: $\beta_{pj} = \gamma_{p0} + \sum_{q=1}^Q \gamma_{pq} Z_{qj} + U_{pj}$

- $\epsilon \sim N(0, \sigma^2)$

- $U_j \sim MVN(0, \boldsymbol{\tau})$


$$\begin{bmatrix} \tau_{00} & \\ \tau_{01} & \tau_{11} \end{bmatrix}$$

Repeated Measures

- Our example thus far has been cross-sectional, but sometimes data are repeated measures, with multiple responses on the outcome per person
- For example, students in a driver's ed class are shown 51 pictures, and they change slightly while they're looking
 - Outcome variable is log reaction time to notice a change
- Here, trials/measures are nested within students
- Level 1 variables are characteristics of the pictures
 - How big is the change? Is the picture related to driving? Etc.
- Level 2 variables are characteristics of the person
 - Gender? Age? etc.

Coding an MLM in R

- Data and example from Hoffman and Rovine (2007)
- Outcome of interest is log of reaction time for participants to detect a change to a picture
- Pictures varied on how salient the change in the picture was (salience)
- One of the primary questions was how age related to reaction time, given picture characteristics

Example Equations: Repeated Measures

Outcome: log of reaction time to notice difference between pictures (logRT)

Level 1 predictor: salience of change, centered around a constant (c_sal)

Level 2 predictor: age

- Random intercept and random slope
 - Random intercept = how a person's average RT deviates from mean across people, controlling for predictors
 - Random slope = how the relationship between salience and RT deviates from mean across people, controlling for age

Level 1 equation: $\log RT_{ij} = \beta_{0j} + \beta_{1j}c_sal_{ij} + \epsilon_{ij}$

Level 2 intercept equation: $\beta_{0j} = \gamma_{00} + \gamma_{01}age_j + U_{0j}$

Level 2 slope equation: $\beta_{1j} = \gamma_{10} + U_{1j}$

Total: $\log RT_{ij} = \gamma_{00} + \gamma_{01}age_j + \gamma_{10}c_sal_{ij} + U_{0j} + U_{1j}c_sal_{ij} + \epsilon_{ij}$

Coding an MLM in R

- Repeated measures data, with many pictures per participant

id	sex	age	NAME	rt_sec	Item	meaning	salience	lg_rt	oldage	yrs65	c_mean	c_sal
1	1	20	rt_sec1	4.662	1	3.5	4.0	1.5394445	0	0	0.5	1.0
1	1	20	rt_sec2	6.660	2	0.0	3.0	1.8961195	0	0	-3.0	0.0
1	1	20	rt_sec3	6.602	3	4.0	2.0	1.8873726	0	0	1.0	-1.0
1	1	20	rt_sec4	1.332	4	4.0	4.0	0.2866816	0	0	1.0	1.0
1	1	20	rt_sec5	1.332	5	0.0	5.0	0.2866816	0	0	-3.0	2.0
1	1	20	rt_sec7	1.302	7	3.5	4.5	0.2639015	0	0	0.5	1.5
1	1	20	rt_sec8	2.601	8	5.0	1.5	0.9558960	0	0	2.0	-1.5
1	1	20	rt_sec9	23.287	9	0.0	5.0	3.1478953	0	0	-3.0	2.0
1	1	20	rt_sec10	3.330	10	4.0	3.5	1.2029723	0	0	1.0	0.5

Coding an MLM in R

- Data and example from Hoffman and Rovine (2007)

```
# package install, if needed
install.packages("lme4")

# Load dependency
library(lme4)

# Read data
hoffman2007 <- read.csv("hoffman2007.csv")

# Run model
model <- lmer(logRT ~ 1 + c_sal + age + (1 + c_sallid),
              data = hoffman2007,
              REML = FALSE)

# Print output
summary(model)
```

Interpreting Output

Model input information

Linear mixed model fit by maximum likelihood ['lmerMod']
Formula: logRT ~ 1 + c_sal + age + (c_sal | id)
Data: hoffman2007

Fit information

AIC	BIC	logLik	deviance	df.resid
16404.1	16452.7	-8195.0	16390.1	7639

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.5591	-0.7430	-0.1218	0.6220	4.1749

Random effects

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
id	(Intercept)	0.0224610	0.14987	
	c_sal	0.0004847	0.02201	-0.05
Residual		0.4870854	0.69792	

Number of obs: 7646, groups: id, 153

Fixed effects

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	1.0429887	0.0259398	40.21
c_sal	-0.1518363	0.0075316	-20.16
age	0.0143094	0.0005319	26.90

Fixed Effects

- (Intercept) = γ_{00} = average log reaction time across all people across all photos, controlling for predictors
- c_sal = γ_{10} = average effect of photo salience on reaction time across all people, controlling for age
- age = γ_{01} = effect of age on average log reaction time, controlling for c_sal

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	1.0429887	0.0259398	40.21
c_sal	-0.1518363	0.0075316	-20.16
age	0.0143094	0.0005319	26.90

Random Effects

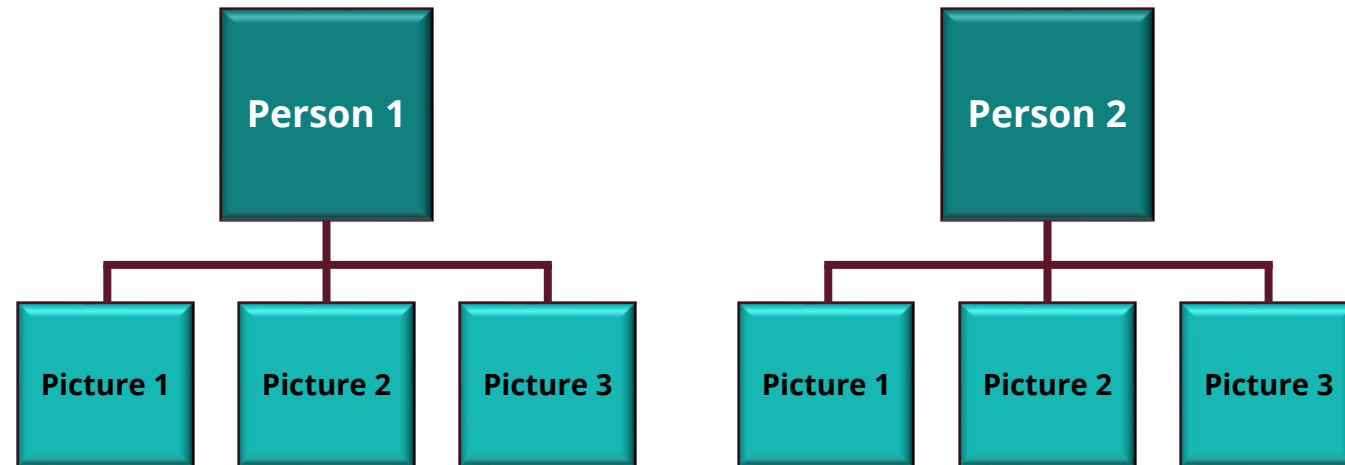
- (Intercept) = $\tau_0^2 = \text{var}(U_{0j})$ = variance describing how people's mean RT vary around grand mean intercept
- c_sal = $\tau_1^2 = \text{var}(U_{1j})$ = variance describing how people's relationship with salience and RT varies around grand mean slope
- Residual = $\sigma^2 = \text{var}(\epsilon_{ij})$ = residual describing how people's responses vary around their own mean

Random effects:

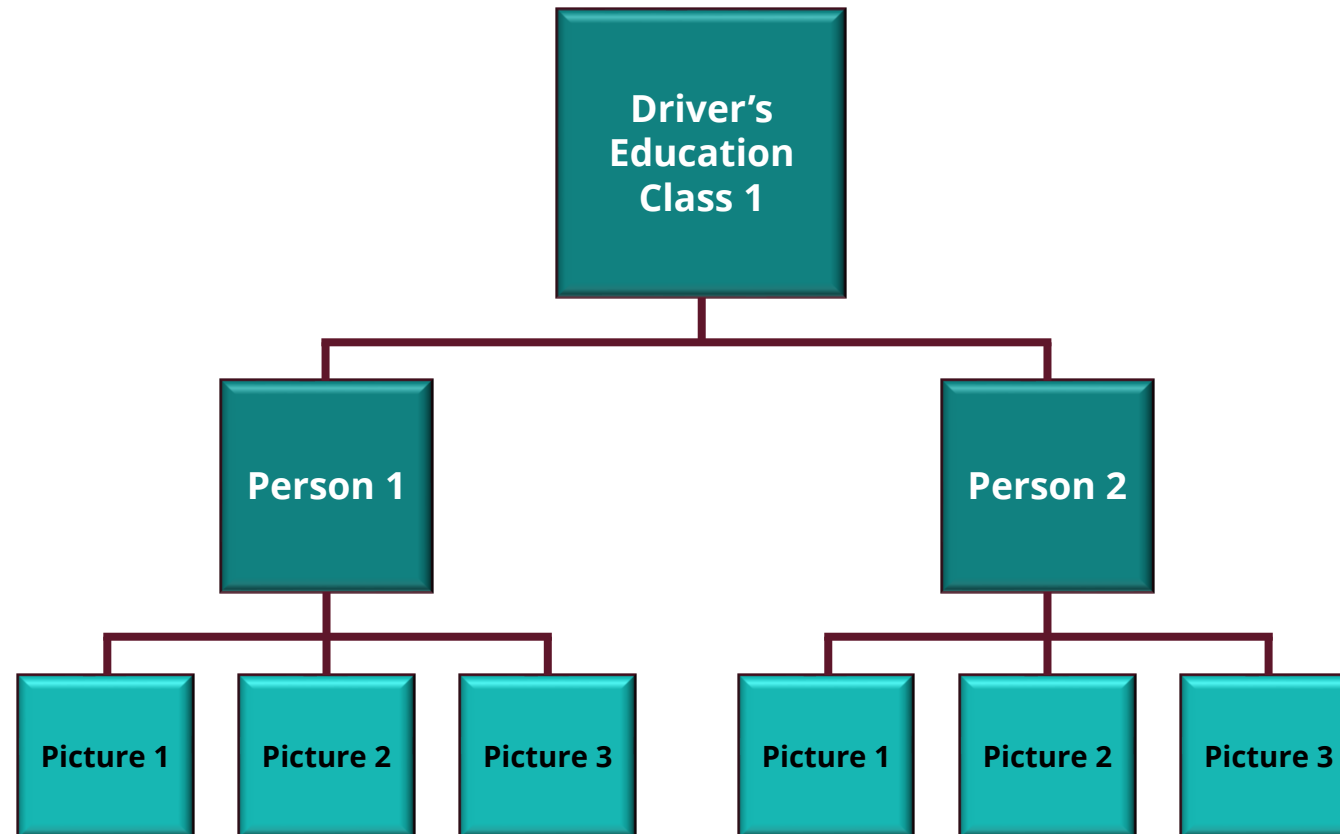
Groups	Name	Variance	Std.Dev.	Corr
id	(Intercept)	0.0224610	0.14987	
	c_sal	0.0004847	0.02201	-0.05
Residual		0.4870854	0.69792	

Number of obs: 7646, groups: id, 153

Two-Level MLM



Three-Level MLM



Three-Level MLM: Equations

Outcome: log of reaction time to notice difference between pictures (logRT)

Intercept-only model: no predictors yet

Level 1 equation: $\log RT_{ijk} = \pi_{0jk} + \epsilon_{ijk}$

Level 2 intercept equation: $\pi_{0jk} = \beta_{00k} + r_{0jk}$

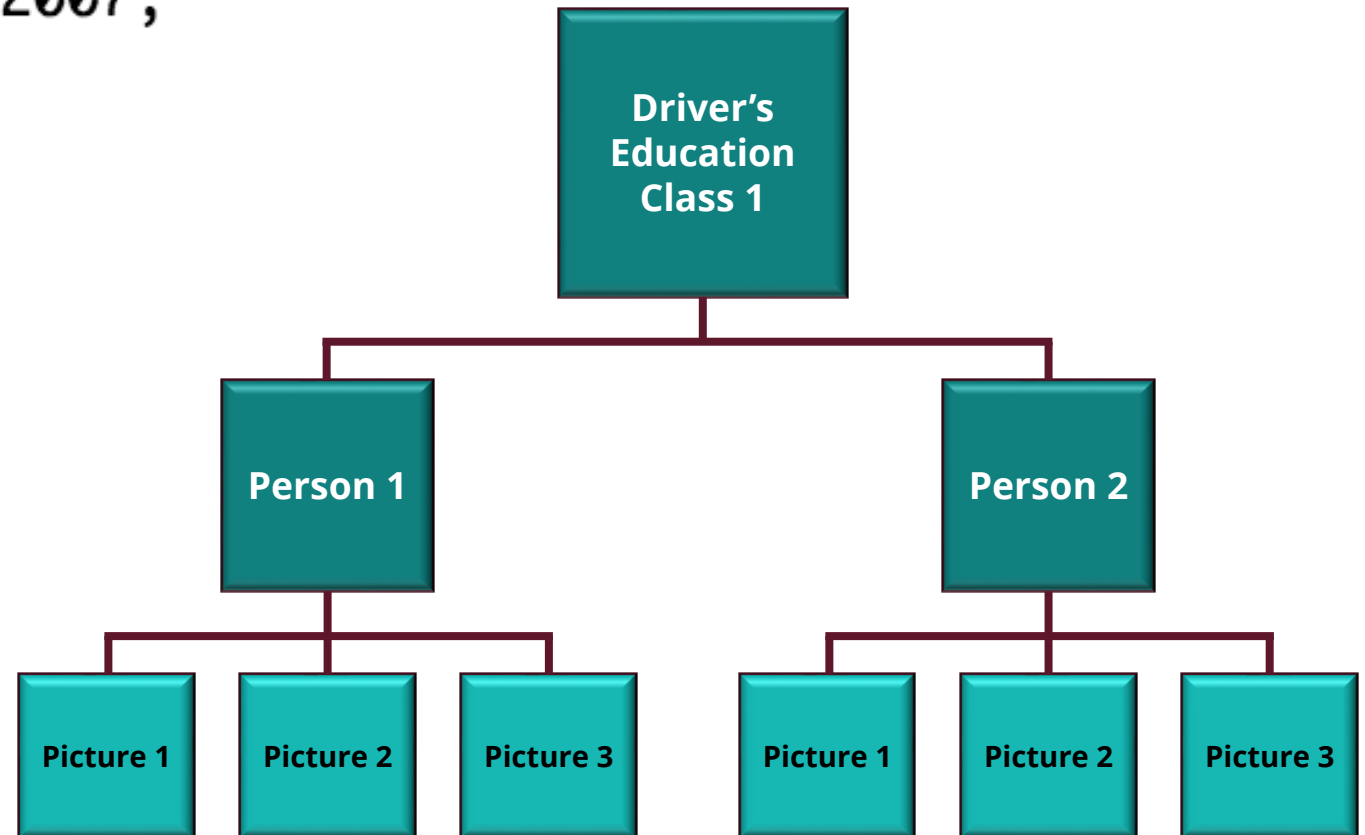
Level 3 intercept equation: $\beta_{00k} = \gamma_{000} + U_{00k}$

Total: $\log RT_{ijk} = \gamma_{000} + U_{00k} + r_{0jk} + \epsilon_{ijk}$

- $\log RT_{ijk}$ is log reaction time for picture i , person j , in class k
- γ_{000} is grand mean RT across all pictures, all people, all classes
- U_{00k} is residual term describing how class k 's average deviates from grand mean
- r_{0jk} is residual term describing how person j deviates from class mean
- ϵ_{ijk} is residual term describing how logRT to picture i deviates from person's mean

Three-Level MLM: Code

```
# three-level null model  
model3 <- lmer(logRT ~ 1 + (1|id) + (1|class),  
               data = hoffman2007,  
               REML = FALSE)  
summary(model3)
```



Three-Level MLM: Output

Random effects:

Groups	Name	Variance	Std.Dev.
id	(Intercept)	0.02556	0.1599
class	(Intercept)	0.14045	0.3748
Residual		0.51573	0.7181

Number of obs: 7646, groups: id, 153; class, 12

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	1.5885	0.1093	12.0212	14.54	0.00000000544 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Long vs Wide Data

- We often work with “wide” data: one row per person

id	meaning1	meaning2	meaning3	salience1	salience2	salience3
1	3.5	0	4	4	3	2
2	3.5	0	4	4	3	2

- To run a repeated measures model, we need “long” data: one row per measurement

id	Item	meaning	salience
1	1	3.5	4
1	2	0.0	3
1	3	4.0	2
2	1	3.5	4
2	2	0.0	3
2	3	4.0	2

- In many cases, this will require transposing data

Summary

- Multilevel modelling is used in data structures where responses are clustered (e.g., students in classes, responses within person)
- It involves partitioning total variance to variance within and between clusters
- We reviewed formulae for cross-sectional, repeated measures, and three-level MLMs
- We implemented and interpreted a repeated measures and three-level MLM in R

Measurement Modelling Overview

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Measurement Modelling Overview

Section Learning Objectives

Define measurement modelling

List the two elements of a structural equation model

Code a confirmatory factor analysis using *lavaan* in R

Interpret loadings, variances, and fit indices for CFA

What are Measurement Models?

- In educational and psychological research, we often work with latent constructs
 - Theoretical entities that account for characteristics or behaviors (Bandalos, 2018)
 - e.g., motivation, depression, anxiety
- We create items (prompts, questions) to capture elements of these constructs so we can theorize about them
- Measurement modelling is the process of relating these items to the latent variables they (supposedly) capture
- Any kind of relationship between an item and a latent variable is a measurement model
 - Sum scores
 - Single items
 - Factor analysis

Structural Equation Modelling

- Combination of a type of measurement modelling (e.g., CFA) and analyzing relationships between variables of interest (path analysis)
- Three inputs, three outputs

In:

1. Theory of causal relationships between variables
2. Specification of relationships between variables of interest
3. Data

Out:

1. Numeric estimates of model parameters for hypothesized relationships between variables
2. Implications of the model not directly specified (e.g., W and Y unrelated, controlling for Z)
3. Fit indices reflecting degree to which testable implications are supported by data

Example CFA

- We'll now look at an example of a confirmatory factor analysis, a very common and useful measurement model
- First, we'll look at the data structure and discuss the models we'll run
- Then, we'll look at two different (equivalent) ways to identify the model and code it
- Finally, we'll look at output and interpret loadings, variances, and fit indices

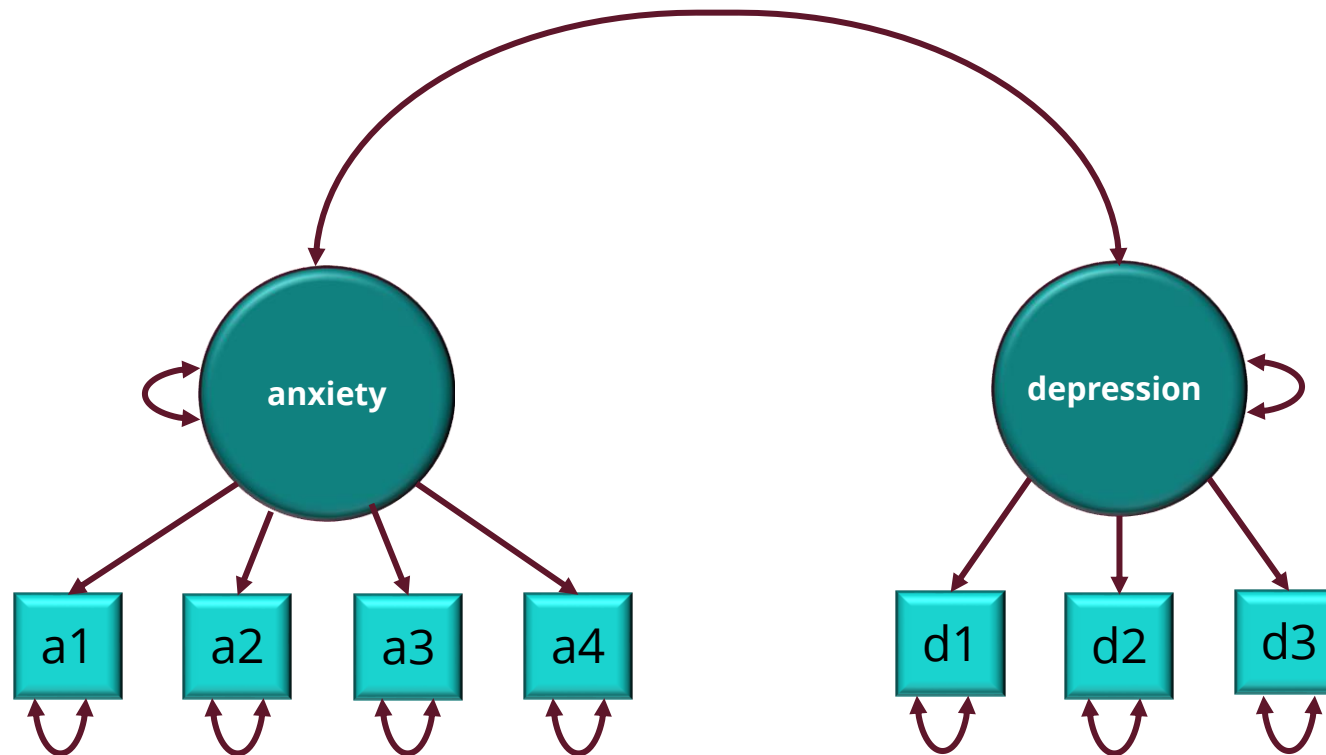
Example CFA: Data Structure

- Simulated data
- Responses range from 1 to 5 for questions about anxiety and depression symptoms

a1	a2	a3	a4	d1	d2	d3
1	2	1	1	3	2	2
1	2	1	2	2	1	2
3	1	1	1	1	1	2
3	1	1	4	2	2	1
1	2	1	2	3	1	3
1	1	2	2	2	2	1
5	2	2	2	2	1	2
2	1	1	1	1	1	2

Example CFA: Model Specification

- Two-factor model: anxiety and depression
- Four items assessing anxiety: a1 - a4
- Three items assessing depression: d1 - d3



Example CFA: Code

```
# 0 Load dependencies and data -----  
  
library(lavaan)  
data <- read.csv("sim_cfa.csv")  
  
# 1 CFA -----  
  
model <- '  
  # Item loadings  
  ANX =~ NA*a1 + a2 + a3 + a4  
  DEP =~ NA*d1 + d2 + d3  
  
  # Factor co/variances  
  ANX ~~ 1*ANX  
  DEP ~~ 1*DEP  
  ANX ~~ DEP  
'  
  
model_cfa <- cfa(model, data)  
  
summary(model_cfa, fit.measures = TRUE, standardized = TRUE)
```

Example CFA: Loadings and Variances

Latent Variables:

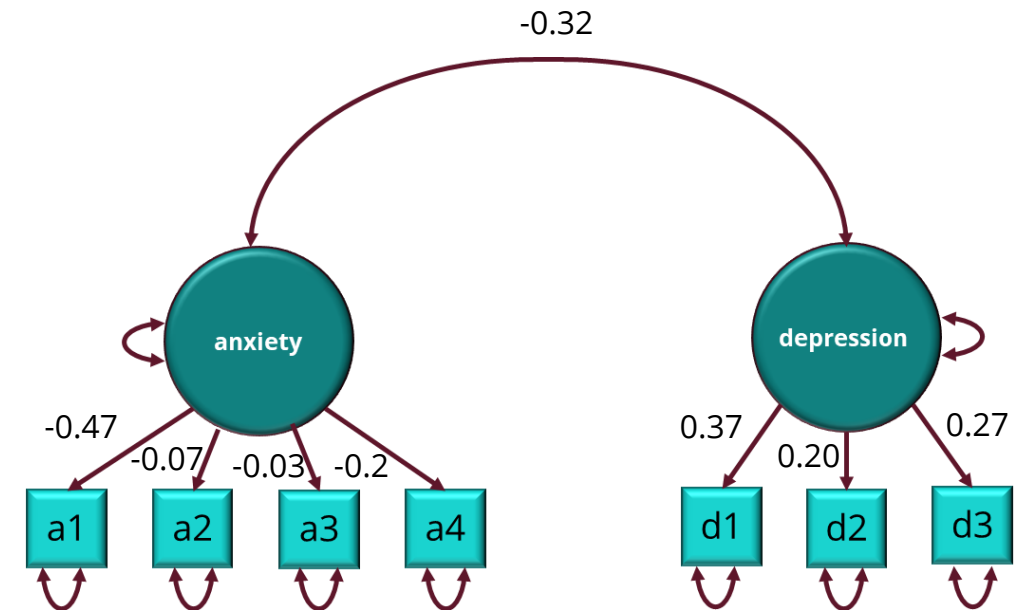
	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
ANX =~						
a1	-0.467	0.140	-3.345	0.001	-0.467	-0.549
a2	-0.073	0.040	-1.800	0.072	-0.073	-0.112
a3	-0.026	0.038	-0.691	0.489	-0.026	-0.041
a4	-0.200	0.064	-3.116	0.002	-0.200	-0.287
DEP =~						
d1	0.370	0.072	5.106	0.000	0.370	0.454
d2	0.195	0.047	4.191	0.000	0.195	0.273
d3	0.274	0.057	4.839	0.000	0.274	0.371

Covariances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
ANX ~~						
DEP	-0.324	0.120	-2.693	0.007	-0.324	-0.324

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
ANX	1.000				1.000	1.000
DEP	1.000				1.000	1.000
.a1	0.507	0.130	3.888	0.000	0.507	0.699
.a2	0.418	0.021	19.489	0.000	0.418	0.988
.a3	0.415	0.021	19.940	0.000	0.415	0.998
.a4	0.445	0.033	13.680	0.000	0.445	0.918
.d1	0.528	0.056	9.393	0.000	0.528	0.794
.d2	0.474	0.028	16.779	0.000	0.474	0.926
.d3	0.469	0.036	12.903	0.000	0.469	0.862



Example CFA: Fit Indices

Model Test User Model:

Test statistic	10.084
Degrees of freedom	13
P-value (Chi-square)	0.687

Model Test Baseline Model:

Test statistic	84.604
Degrees of freedom	21
P-value	0.000

User Model versus Baseline Model:

Comparative Fit Index (CFI)	1.000
Tucker-Lewis Index (TLI)	1.074

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-6121.951
Loglikelihood unrestricted model (H1)	-6116.909
Akaike (AIC)	12273.902
Bayesian (BIC)	12344.171
Sample-size adjusted Bayesian (BIC)	12296.538

Root Mean Square Error of Approximation:

RMSEA	0.000
90 Percent confidence interval - lower	0.000
90 Percent confidence interval - upper	0.028
P-value RMSEA ≤ 0.05	1.000

Standardized Root Mean Square Residual:

SRMR	0.021
------	-------

Example CFA: Code

```
# Unit loading identification
```

```
model <- '  
  # Item loadings  
  ANX =~ a1 + a2 + a3 + a4  
  DEP =~ d1 + d2 + d3  
  
  # Factor co/variances  
  ANX ~~ ANX  
  DEP ~~ DEP  
  ANX ~~ DEP  
'
```

```
model_cfa <- cfa(model, data)
```

```
summary(model_cfa, fit.measures = TRUE, standardized = TRUE)
```

Example CFA: Loadings and Variances

Latent Variables:

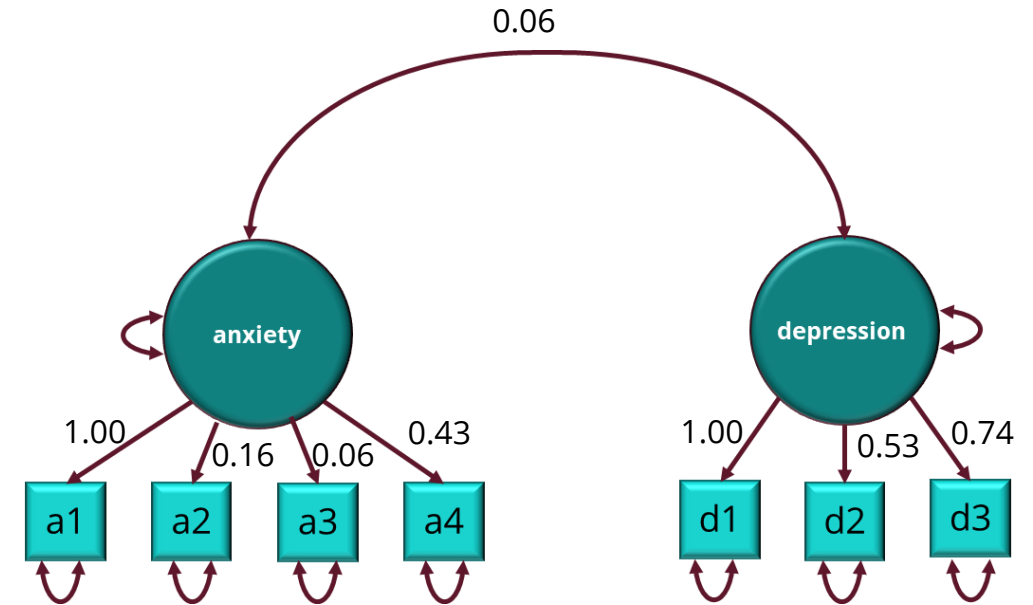
	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
ANX =~						
a1	1.000				0.467	0.549
a2	0.155	0.110	1.416	0.157	0.073	0.112
a3	0.056	0.085	0.662	0.508	0.026	0.041
a4	0.428	0.247	1.732	0.083	0.200	0.287
DEP =~						
d1	1.000				0.370	0.454
d2	0.528	0.180	2.932	0.003	0.195	0.273
d3	0.741	0.255	2.906	0.004	0.274	0.371

Covariances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
ANX ~~						
DEP	0.056	0.021	2.691	0.007	0.324	0.324

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
ANX	0.218	0.130	1.673	0.094	1.000	1.000
DEP	0.137	0.054	2.553	0.011	1.000	1.000
.a1	0.507	0.130	3.888	0.000	0.507	0.699
.a2	0.418	0.021	19.489	0.000	0.418	0.988
.a3	0.415	0.021	19.940	0.000	0.415	0.998
.a4	0.445	0.033	13.680	0.000	0.445	0.918
.d1	0.528	0.056	9.393	0.000	0.528	0.794
.d2	0.474	0.028	16.779	0.000	0.474	0.926
.d3	0.469	0.036	12.903	0.000	0.469	0.862



From CFA to SEM

- Once you have your measurement model sorted (in our case, the CFA), you can move on to SEM
- You can add more variables and use the latent variables from the measurement model as predictors or outcomes for other latent variables or observed variables
- Our focus is on the measurement model, so we won't cover SEM further

Summary

- Measurement modelling is the process of relating items to latent constructs
- Confirmatory factor analysis is one kind of measurement model
- Structural equation modelling combines measurement modelling (e.g., CFA) with structural modelling of relationships between latent variables
- We implemented a confirmatory factor analysis using lavaan in R

Multilevel Measurement Modelling

A series of vertical bars of varying heights and colors (maroon and gold) on the left side of the slide.

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A large maroon circle with a gold number 3 in the center.

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Multilevel Measurement Modelling

Section Learning Objectives

Define a multilevel measurement model (MLMM)

Understand when to use an MLMM

Describe two different approaches to multilevel measurement modelling

State two issues that arise from not using an MLMM

What is a Multilevel Measurement Model?

So far, we've introduced and reviewed two modelling frameworks

1. Multilevel modelling for clustered data
2. Measurement modelling in single-level data

We use multilevel measurement models (MLMMs) when these circumstances collide, i.e. when:

- We want to conduct measurement modelling (linking items to factors)
- In clustered data structures

Why use an MLMM?

MLMMs are a mix of measurement modelling and clustered data modelling

If you don't use one, you run into two main issues:

No Measurement Modelling

no accounting for measurement error

No Multilevel Modelling

no accounting for clustered data

Dyer, N., Hanges, P., & Hall, R. (2005). Applying multilevel confirmatory factor analysis techniques to the study of leadership. *The Leadership Quarterly*, 16, 149-167. doi:10.1016/j.leaqua.2004.09.009

Huang, F. (2017). *Conducting Multilevel Confirmatory Factor Analysis Using R*.

Muthén, B. O. (1994). Multilevel Covariance Structure Analysis. *Sociological Methods & Research*, 22(3), 376-398. doi:10.1177/0049124194022003006

No Measurement Modelling

- Assumes perfect measurement
- In our MLM example, we used hours studied and teacher salary, which could plausibly have no measurement error
- In our SEM example, we talked about anxiety and depression which are more clearly latent constructs

Why use an MLMM?

MLMMs are a mix of measurement modelling and clustered data modelling

If you don't use one, you run into two main issues:

No Measurement Modelling

no accounting for measurement error

No Multilevel Modelling

no accounting for clustered data

Dyer, N., Hanges, P., & Hall, R. (2005). Applying multilevel confirmatory factor analysis techniques to the study of leadership. *The Leadership Quarterly*, 16, 149-167. doi:10.1016/j.leaqua.2004.09.009

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Muthén, B. O. (1994). Multilevel Covariance Structure Analysis. *Sociological Methods & Research*, 22(3), 376-398. doi:10.1177/0049124194022003006

No Multilevel Modelling

- Without MLM, we don't account for clustered data
- Violates assumption of independence
- Can bias parameter estimates, standard errors, and model fit
- Denies ability to ask multilevel questions
 - Some constructs have different factor structures across levels
 - Some constructs only exist at the group level, e.g., “classroom environment”

MLM and SEM: Analogous Frameworks

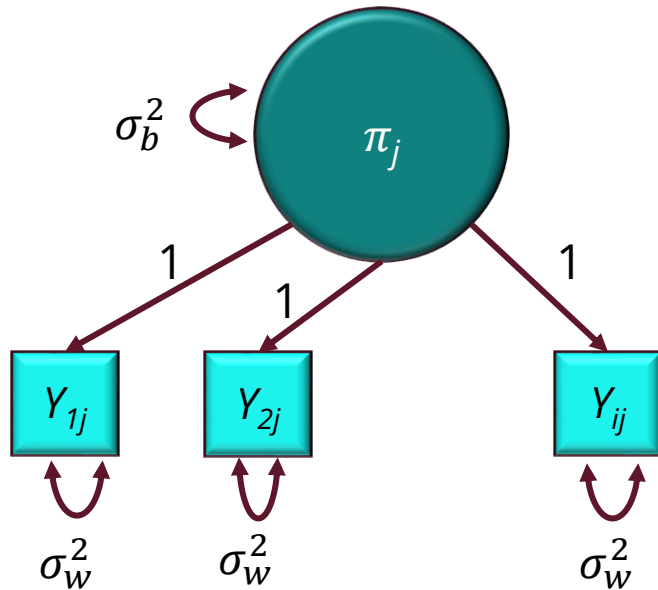
MLM

$$Y_{ij} = \beta + U_j + \epsilon_{ij}$$

SEM

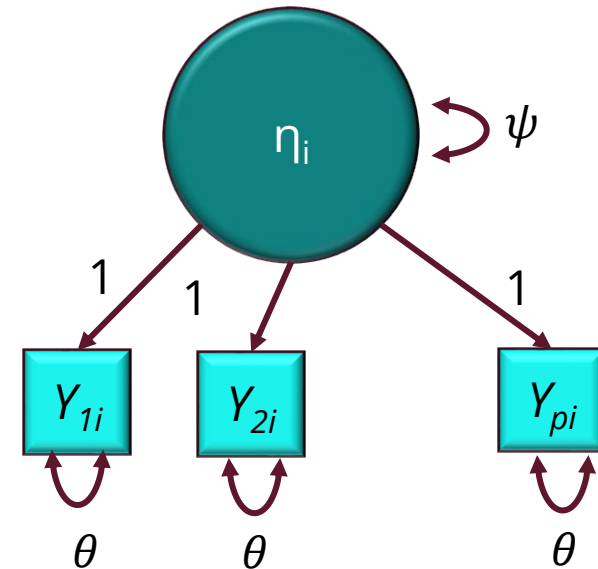
$$Y_{pi} = v_p + \lambda_{pq}\eta_{qi} + \epsilon_{pi}$$

Between variance



Within variance

Common variance



Unique variance

Two Frameworks: MLM or SEM

We can adapt each of the two frameworks reviewed to MLMM

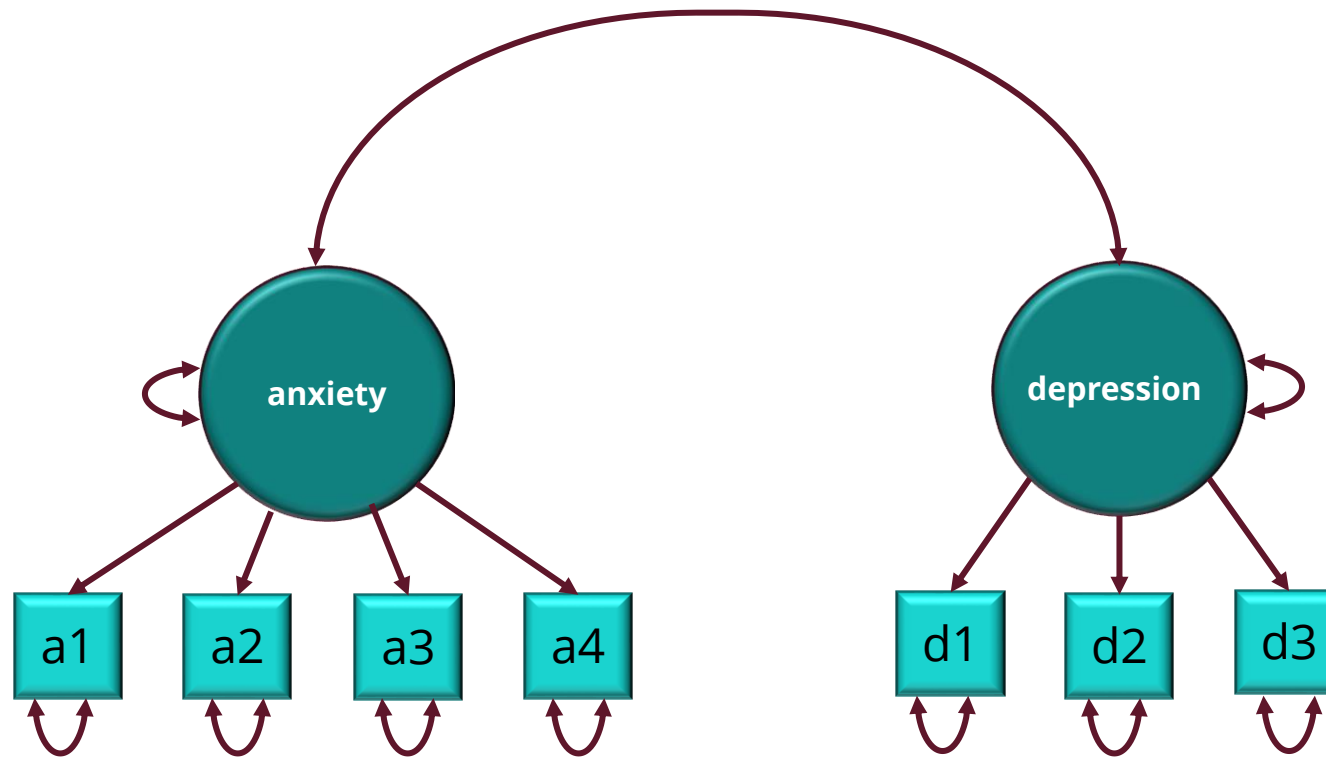
MLM: Use multilevel modelling for measurement modelling



Two Frameworks: MLM or SEM

We can adapt each of the two frameworks reviewed to MLMM

SEM: Add multilevel modelling to measurement modelling



Unclustered Factor Model: MLM vs SEM

- Let's run a factor model in each framework assuming no clustering, to compare and set the stage for adding another level
- We're not getting into data preparation (that's in the next section) or model details, just looking at code and output

Example: Teacher perception of student engagement

- One factor
- Six items
 - e.g., "Students generally like this school."
- Data from: Huang F.L. & Cornell D.G. (2015). Factor structure of the high school teacher version of the authoritative school climate survey. *Journal of Psychoeducational Assessment*, 1557-5144. doi:10.1177/0734282915621439.

Unclustered Factor Model: MLM vs SEM

MLM data structure

	response	x1	x2	x3	x4	x5	x6	teacherID
1	4	1	0	0	0	0	0	1
2	4	0	1	0	0	0	0	1
3	3	0	0	1	0	0	0	1
4	4	0	0	0	1	0	0	1
5	4	0	0	0	0	1	0	1
6	3	0	0	0	0	0	1	1
7	6	1	0	0	0	0	0	2
8	6	0	1	0	0	0	0	2

```
data_t <- data %>%
  rename(schoolID = sid) %>%
  mutate(teacherID = 1:nrow(data)) %>% # add row for teacher ID
  pivot_longer(
    cols = starts_with("x"),
    names_to = "item",
    names_prefix = "x",
    values_to = "response"
  ) %>%
  mutate(
    x1 = ifelse(item == 1, 1, 0),
    x2 = ifelse(item == 2, 1, 0),
    x3 = ifelse(item == 3, 1, 0),
    x4 = ifelse(item == 4, 1, 0),
    x5 = ifelse(item == 5, 1, 0),
    x6 = ifelse(item == 6, 1, 0)
  ) %>%
  select(response, x1:x6, item, teacherID, schoolID)
```

SEM data structure

	x1	x2	x3	x4	x5	x6	sid
1	4	4	3	4	4	3	55
2	6	6	4	6	4	4	55
3	5	5	2	6	4	4	55
4	4	4	5	5	3	3	55
5	4	4	4	4	3	2	55
6	4	4	2	4	2	2	55
7	5	4	3	5	3	3	41
8	3	2	2	3	3	2	41

Unclustered Factor Model: MLM vs SEM

MLM code: lme4

```
mlm <- lmer(response ~ 1 + x2 + x3 + x4 + x5 + x6 + (1|teacherID),  
            data = data_t,  
            REML = TRUE)  
summary(mlm)
```

SEM code

```
sem <- '  
    engagement =~ NA*x1 + x2 + x3 + x4 + x5 + x6  
  
    engagement ~~ 1*engagement  
'  
sem_fit <- cfa(model = sem, data = data, meanstructure = TRUE)  
summary(sem_fit,  
        fit.measures = TRUE,  
        standardized = TRUE)
```

Unclustered Factor Model: MLM vs SEM

MLM results: lme4

Random effects:

Groups	Name	Variance	Std.Dev.
	teacherID (Intercept)	0.5620	0.7497
	Residual	0.5496	0.7414
Number of obs: 23364, groups: teacherID, 3894			

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	4.7774	0.0169	282.75
x2	-0.1721	0.0168	-10.24
x3	-1.2360	0.0168	-73.56
x4	-0.3983	0.0168	-23.71
x5	-0.5354	0.0168	-31.87
x6	-0.8318	0.0168	-49.51

SEM results

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement =~						
x1	0.736	0.012	62.120	0.000	0.736	0.840
x2	0.840	0.013	63.372	0.000	0.840	0.851
x3	0.700	0.018	39.239	0.000	0.700	0.599
x4	0.550	0.016	33.425	0.000	0.550	0.524
x5	0.804	0.016	50.571	0.000	0.804	0.728
x6	0.821	0.016	51.505	0.000	0.821	0.738

Intercepts:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
.x1	4.777	0.014	340.099	0.000	4.777	5.450
.x2	4.605	0.016	291.129	0.000	4.605	4.665
.x3	3.541	0.019	189.049	0.000	3.541	3.030
.x4	4.379	0.017	260.388	0.000	4.379	4.173
.x5	4.242	0.018	239.665	0.000	4.242	3.841
.x6	3.946	0.018	221.289	0.000	3.946	3.546
engagement	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement	1.000				1.000	1.000
.x1	0.226	0.007	30.699	0.000	0.226	0.294
.x2	0.268	0.009	29.443	0.000	0.268	0.275
.x3	0.876	0.021	41.111	0.000	0.876	0.641
.x4	0.799	0.019	42.095	0.000	0.799	0.725
.x5	0.573	0.015	37.956	0.000	0.573	0.469
.x6	0.564	0.015	37.582	0.000	0.564	0.455

MLM: nlme approach

SEM code

```
sem <- '  
  
    engagement =~ NA*x1 + x2 + x3 + x4 + x5 + x6  
  
    engagement ~~ 1*engagement  
  
'  
sem_fit <- cfa(model = sem, data = data, meanstructure = TRUE)  
summary(sem_fit,  
        fit.measures = TRUE,  
        standardized = TRUE)
```

MLM code: nlme

```
mlm_nlme <- lme(  
  response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6,  
  data = data_t,  
  random = ~ 1 | teacherID,  
  weights = varIdent(form = ~ 1 | item)  
)  
summary(mlm_nlme)  
  
# level-1 variance components  
(c(1.000, coef(mlm_nlme$modelStruct$varStruct, unconstrained = FALSE))*mlm_nlme$sigma)^2  
  
# level-2 error variance (factor variance)  
.7604**2
```


Unclustered Factor Model: MLM vs SEM

MLM Results: nlme

Random effects:

Formula: ~1 | teacherID
(Intercept) Residual
StdDev: 0.7603665 0.4741239

Variance function:

Structure: Different standard deviations per stratum

Formula: ~1 | item

Parameter estimates:

	1	2	3	4	5	6
1.000000	1.144862	1.967339	1.893319	1.607276	1.600786	

Fixed effects: response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6

	Value	Std.Error	DF	t-value	p-value
x1	4.777350	0.01435974	19465	332.6907	0
x2	4.605290	0.01497126	19465	307.6087	0
x3	3.541346	0.01928487	19465	183.6334	0
x4	4.379045	0.01885230	19465	232.2817	0
x5	4.241911	0.01725123	19465	245.8903	0
x6	3.945557	0.01721636	19465	229.1749	0

Level 1 error variances rescaled:

	2	3	4	5	6
0.2247934	0.2946390	0.8700454	0.8058071	0.5807167	0.5760366

SEM results

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement =~						
x1	0.736	0.012	62.120	0.000	0.736	0.840
x2	0.840	0.013	63.372	0.000	0.840	0.851
x3	0.700	0.018	39.239	0.000	0.700	0.599
x4	0.550	0.016	33.425	0.000	0.550	0.524
x5	0.804	0.016	50.571	0.000	0.804	0.728
x6	0.821	0.016	51.505	0.000	0.821	0.738

Intercepts:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
.x1	4.777	0.014	340.099	0.000	4.777	5.450
.x2	4.605	0.016	291.129	0.000	4.605	4.665
.x3	3.541	0.019	189.049	0.000	3.541	3.030
.x4	4.379	0.017	260.388	0.000	4.379	4.173
.x5	4.242	0.018	239.665	0.000	4.242	3.841
.x6	3.946	0.018	221.289	0.000	3.946	3.546
engagement	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement	1.000				1.000	1.000
.x1	0.226	0.007	30.699	0.000	0.226	0.294
.x2	0.268	0.009	29.443	0.000	0.268	0.275
.x3	0.876	0.021	41.111	0.000	0.876	0.641
.x4	0.799	0.019	42.095	0.000	0.799	0.725
.x5	0.573	0.015	37.956	0.000	0.573	0.469
.x6	0.564	0.015	37.582	0.000	0.564	0.455

MLM: nlme approach

SEM code

```
# SEM with loadings constrained to 1
sem <- '

    engagement =~ 1*x1 + 1*x2 + 1*x3 + 1*x4 + 1*x5 + 1*x6

    engagement ~~ engagement

'

sem_fit <- cfa(model = sem, data = data, meanstructure = TRUE)
summary(sem_fit,
        fit.measures = TRUE,
        standardized = TRUE)
```

MLM code: nlme

```
mlm_nlme <- lme(
  response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6,
  data = data_t,
  random = ~ 1 | teacherID,
  weights = varIdent(form = ~ 1 | item)
)
summary(mlm_nlme)
```

```
# level-1 variance components
```

```
(c(1.000, coef(mlm_nlme$modelStruct$varStruct, unconstrained = FALSE))*mlm_nlme$sigma)^2
```

```
# level-2 error variance (factor variance)
```

```
.7604**2
```

Latent Variable Variance

MLM Results: nlme

Random effects:

Formula: ~1 | teacherID
(Intercept) Residual
StdDev: 0.7603665 0.4741239

Variance function:

Structure: Different standard deviations per stratum

Formula: ~1 | item

Parameter estimates:

	1	2	3	4	5	6
	1.000000	1.144862	1.967339	1.893319	1.607276	1.600786
Fixed effects: response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6						
	Value	Std.Error	DF	t-value	p-value	
x1	4.777350	0.01435974	19465	332.6907	0	
x2	4.605290	0.01497126	19465	307.6087	0	
x3	3.541346	0.01928487	19465	183.6334	0	
x4	4.379045	0.01885230	19465	232.2817	0	
x5	4.241911	0.01725123	19465	245.8903	0	
x6	3.945557	0.01721636	19465	229.1749	0	

Level 2 error variance:

SEM results

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement =~						
x1	1.000				0.760	0.849
x2	1.000				0.760	0.814
x3	1.000				0.760	0.632
x4	1.000				0.760	0.646
x5	1.000				0.760	0.706
x6	1.000				0.760	0.708

Intercepts:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
.x1	4.777	0.014	332.734	0.000	4.777	5.332
.x2	4.605	0.015	307.647	0.000	4.605	4.930
.x3	3.541	0.019	183.655	0.000	3.541	2.943
.x4	4.379	0.019	232.311	0.000	4.379	3.723
.x5	4.242	0.017	245.920	0.000	4.242	3.941
.x6	3.946	0.017	229.202	0.000	3.946	3.673
engagement	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement	0.578	0.015	39.098	0.000	1.000	1.000
.x1	0.225	0.007	31.058	0.000	0.225	0.280
.x2	0.295	0.009	34.130	0.000	0.295	0.338
.x3	0.870	0.021	40.802	0.000	0.870	0.601
.x4	0.806	0.020	40.535	0.000	0.806	0.582
.x5	0.581	0.015	39.121	0.000	0.581	0.501
.x6	0.576	0.015	39.079	0.000	0.576	0.499

Summary

- A multilevel measurement model (MLMM) is used to do measurement modelling in clustered data
- We outlined two approaches: (1) using an MLM framework and (2) using an SEM framework
- We ran an unclustered (i.e., single-level) measurement model in each framework to demonstrate the analogy and preview the similarities and differences

Multilevel Measurement Models in a Multilevel Modelling Framework

A series of vertical bars of varying heights and colors (maroon, gold, and light purple) on the left side of the slide.

4

A large maroon circle with a gold number 4 in the center.

4

MLMMs in MLM Framework

Section Learning Objectives

State the 5-step process for conducting an MLMM in MLM framework

Restructure data to be used with MLM framework

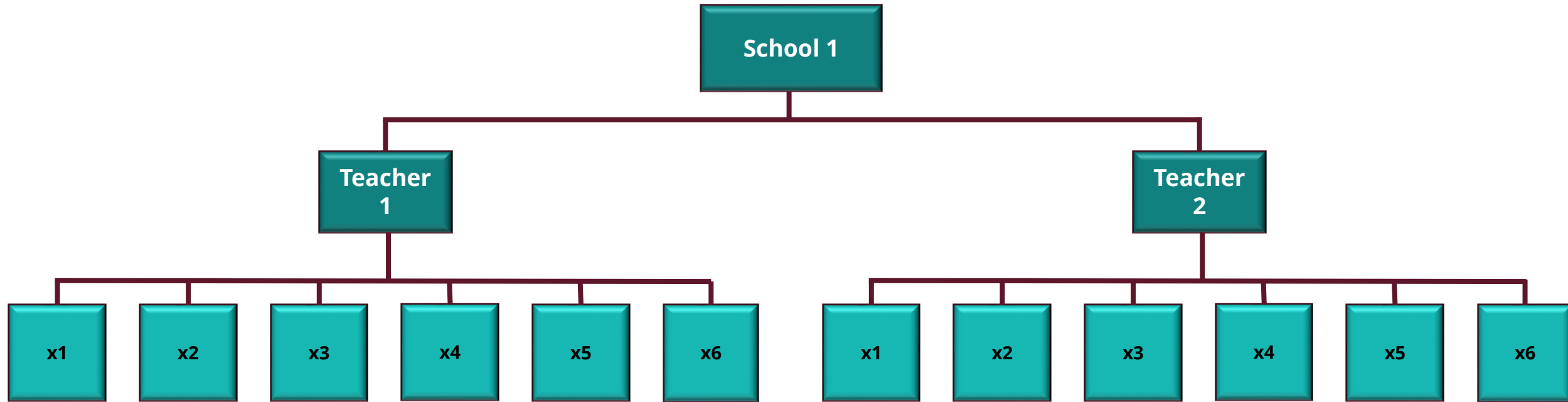
List and interpret equations for multilevel measurement models in MLM framework

Execute multilevel measurement model in R package *nlme* and interpret output

Our Example

- Data structure: teachers within schools
 - Between 5 and 50 teachers per school
 - 254 schools
- Outcome: perception of student engagement
- Outcome variance can be decomposed into three parts
 - Within teachers: variance across item responses
 - Between teachers: variance across teachers within schools
 - Between schools: variance across schools
- 6 items assessing student engagement
 - 1 = strongly disagree, 2 = disagree, 3 = somewhat disagree, 4 = somewhat agree, 5 = agree, 6 = strongly agree
- Random subset of data from Huang and Cornell (2015)
 - <https://github.com/flh3/pubdata/blob/main/MLCFA/raw.csv>

Our Example



- x1: Students generally like this school.
- x2: Students are proud to be at this school.
- x3: Students finish their homework at this school.
- x4: Students hate going to school. (reverse-coded)
- x5: Getting good grades is very important to most students here.
- x6: Most students want to learn as much as they can at this school.

MLMM: Equations

- **Outcome:** response to item
- **Level 1 equation:** $Y_{ijk} = \pi_{0jk} + \pi_{1jk}X1_{ijk} + \cdots + \pi_{5jk}X6_{ijk} + \epsilon_{ijk}$
- **Level 2 equations:** $\pi_{0jk} = \beta_{00k} + r_{0jk}$
 $\pi_{1jk} = \beta_{10k}$
...
 $\pi_{5jk} = \beta_{50k}$
- **Level 3 equation:** $\beta_{00k} = \gamma_{000} + U_{00k}$
 $\beta_{10k} = \gamma_{100}$
...
 $\beta_{50k} = \gamma_{500}$
- **Total:** $Y_{ijk} = \gamma_{100}X1_{ijk} + \cdots + \gamma_{500}X6_{ijk} + U_{00k} + r_{0jk} + \epsilon_{ijk}$
 - γ_{000} set to zero

MLMM: What the Model Means

- We are specifying a measurement model for the items, providing a student engagement factor score for each teacher and for each school
- Model results:
 - Intercept for each item that we assume holds across all teachers and schools (testable assumption)
 - Variance across teachers in student engagement factor scores
 - Variance across schools in student engagement factor scores
 - Factor loadings for each item are assumed to be 1 across teachers and schools

Steps for Running MLMM

Step 1

Load data

Step 2

Prepare data for MLMM

Step 3

Check extent of clustering

Step 4

MLMM

Step 5

Interpret output

Step 1

Load data

```
library(RCurl) # for extracting data from web  
library(dplyr) # for data manipulation  
library(tidyr) # for pivoting data  
library(lme4) # for MLMs  
library(performance) # for ICCs
```

```
data <- read.csv(text = getURL("https://raw.githubusercontent.com/flh3/pubdata/main/MLCFA/raw.csv"))
```

Step 2

Prepare data for MLMM

- Our structure is responses to items, clustered within teachers, clustered within schools
- We need long data when working with repeated measures
- That means transposing our dataset
- Our model outcome will be the response to a given item
- Each row is one item, with dummy coded variables

Step 2

Prepare data for MLMM

Before

	x1	x2	x3	x4	x5	x6	schoolID
1	4	4	3	4	4	3	55
2	6	6	4	6	4	4	55
3	5	5	2	6	4	4	55
4	4	4	5	5	3	3	55
5	4	4	4	4	3	2	55
6	4	4	2	4	2	2	55
7	5	4	3	5	3	3	41
8	3	2	2	3	3	2	41
9	4	4	2	4	1	1	41
10	4	4	2	5	2	2	41

After

	response	x1	x2	x3	x4	x5	x6	teacherID	schoolID
1	4	1	0	0	0	0	0	1	55
2	4	0	1	0	0	0	0	1	55
3	3	0	0	1	0	0	0	1	55
4	4	0	0	0	1	0	0	1	55
5	4	0	0	0	0	1	0	1	55
6	3	0	0	0	0	0	1	1	55
7	6	1	0	0	0	0	0	2	55
8	6	0	1	0	0	0	0	2	55
9	4	0	0	1	0	0	0	2	55
10	6	0	0	0	1	0	0	2	55

Step 2

Prepare data for MLMM

Code

```
data_t <- data %>%
  rename(schoolID = sid) %>%
  mutate(teacherID = 1:nrow(data)) %>% # add row for teacher ID
  pivot_longer(
    cols = starts_with("x"),
    names_to = "item",
    names_prefix = "x",
    values_to = "response"
  ) %>%
  mutate(
    x1 = ifelse(item == 1, 1, 0),
    x2 = ifelse(item == 2, 1, 0),
    x3 = ifelse(item == 3, 1, 0),
    x4 = ifelse(item == 4, 1, 0),
    x5 = ifelse(item == 5, 1, 0),
    x6 = ifelse(item == 6, 1, 0)
  ) %>%
  select(response, x1:x6, item, teacherID, schoolID)
```

Step 3

Check extent of clustering

- We check the extent of clustering to ascertain whether we need to go through this multi-step process
- We do this with the intraclass correlation coefficient (ICC), a ratio of variance between clusters to total variance
- With three levels, you can consider variance explained by clustering at level 2:

$$ICC_{L2} = \frac{\tau_0^2}{\phi_0^2 + \tau_0^2 + \sigma^2}$$

- Or variance explained by clustering at level 3:

$$ICC_{L3} = \frac{\phi_0^2}{\phi_0^2 + \tau_0^2 + \sigma^2}$$

- Ranges from 0 to 1, with higher numbers indicating that a larger proportion of item response variance is between clusters
- If the data aren't very clustered (ICC below .05), may not be necessary to take the extra steps accounting for clustering

Step 3

Check degree of clustering

nlme: Unconditional ICCs

```
null_nlme <- lme(  
  response ~ 1,  
  data = data_t,  
  method = "REML",  
  random = ~ 1 | schoolID/teacherID/item,  
  control = lmeControl(optimizer = "Nelder-Mead")  
)  
summary(null_nlme)
```

nlme: Conditional ICCs

```
# conditional  
conditional_nlme <- lme(  
  response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6,  
  data = data_t,  
  method = "REML",  
  random = ~ 1 | schoolID/teacherID/item,  
  control = lmeControl(optimizer = "Nelder-Mead")  
)  
summary(conditional_nlme)
```

Step 3

Check degree of clustering

- nlme outputs standard deviations, so we need to square them to get variances
- item %in% teacherID %in% school ID indicates level-1 variance, σ^2
- teacherID %in% schoolID indicates level-2 variance, τ_0^2
- schoolID indicates level-3 variance, ϕ_0^2

Unconditional

Random effects:

Formula: ~1 | schoolID
(Intercept)

StdDev: 0.4075208

Formula: ~1 | teacherID %in% schoolID
(Intercept)

StdDev: 0.60208

Formula: ~1 | item %in% teacherID %in% schoolID
(Intercept) Residual

StdDev: 0.8006477 0.3341296

Conditional

Random effects:

Formula: ~1 | schoolID
(Intercept)

StdDev: 0.4181061

Formula: ~1 | teacherID %in% schoolID
(Intercept)

StdDev: 0.6386338

Formula: ~1 | item %in% teacherID %in% schoolID
(Intercept) Residual

StdDev: 0.3570088 0.3177386

Step 3

Check degree of clustering

- $ICC_{L2} = \frac{\tau_0^2}{\phi_0^2 + \tau_0^2 + \sigma^2}$

- *Unconditional* $= \frac{0.60^2}{0.60^2 + 0.41^2 + 0.80^2} = 0.30$

- *Conditional* $= \frac{0.64^2}{0.64^2 + 0.42^2 + 0.36^2} = 0.57$

Unconditional

Random effects:

Formula: ~1 | schoolID
(Intercept)

StdDev: 0.4075208

Formula: ~1 | teacherID %in% schoolID
(Intercept)

StdDev: 0.60208

Formula: ~1 | item %in% teacherID %in% schoolID
(Intercept) Residual

StdDev: 0.8006477 0.3341296

- $ICC_{L3} = \frac{\phi_0^2}{\phi_0^2 + \tau_0^2 + \sigma^2}$

- *Unconditional* $= \frac{0.41^2}{0.60^2 + 0.41^2 + 0.80^2} = 0.14$

- *Conditional* $= \frac{0.42^2}{0.64^2 + 0.42^2 + 0.36^2} = 0.25$

Conditional

Random effects:

Formula: ~1 | schoolID
(Intercept)

StdDev: 0.4181061

Formula: ~1 | teacherID %in% schoolID
(Intercept)

StdDev: 0.6386338

Formula: ~1 | item %in% teacherID %in% schoolID
(Intercept) Residual

StdDev: 0.3570088 0.3177386

Step 4

MLMM

$$\text{Total: } Y_{ijk} = \gamma_{100}X1_{ijk} + \cdots + \gamma_{500}X6_{ijk} + U_{00k} + r_{0jk} + \epsilon_{ijk}$$

- The MLM framework allows us to model the relationships between items and latent student engagement factor, and obtain standing of teachers and schools on this factor
- Partitions total variance into 3 levels: within-teacher, between-teacher, and between-schools
- Yields fixed effects for items (average responses to items across all teachers and schools) and random effects for U_{00k} (how school mean varies from grand mean), r_{0jk} (how teacher's mean response varies from school mean), and e_{ijk} (how teacher's response to item deviates from their own mean response across items controlling for L1 predictors)

Step 4

MLMM

$$\text{Total: } Y_{ijk} = \gamma_{100}X1_{ijk} + \cdots + \gamma_{500}X6_{ijk} + U_{00k} + r_{0jk} + \epsilon_{ijk}$$

```
mlmm <- lme(  
  response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6,  
  data = data_t,  
  random = ~ 1 | schoolID/teacherID/item,  
  weights = varIdent(form = ~ 1 | item)  
)  
summary(mlmm)
```

Step 5

Interpret output

Random effects:

Formula: ~1 | schoolID
(Intercept)

StdDev: 0.4181061

Formula: ~1 | teacherID %in% schoolID
(Intercept)

StdDev: 0.6386338

Formula: ~1 | item %in% teacherID %in% schoolID
(Intercept) Residual

StdDev: 0.3570088 0.3177386

Variance function:

Structure: Different standard deviations per stratum

Formula: ~1 | item

Parameter estimates:

	1	2	3	4	5	6
	1.000000	1.273703	2.708233	2.593207	2.100547	2.124616

Fixed effects: response ~ 0 + x1 + x2 + x3 + x4 + x5 + x6

	Value	Std.Error	DF	t-value	p-value
x1	4.741659	0.02977883	19465	159.2292	0
x2	4.569599	0.03004852	19465	152.0740	0
x3	3.505655	0.03241930	19465	108.1348	0
x4	4.343354	0.03217453	19465	134.9935	0
x5	4.206220	0.03122896	19465	134.6897	0
x6	3.909866	0.03127114	19465	125.0311	0

$$\text{Total: } Y_{ijk} = \gamma_{100}X1_{ijk} + \cdots + \gamma_{500}X6_{ijk} + U_{00k} + r_{0jk} + \epsilon_{ijk}$$

Step 5

Interpret output

$$\text{Total: } Y_{ijk} = \gamma_{100}X1_{ijk} + \cdots + \gamma_{500}X6_{ijk} + U_{00k} + r_{0jk} + \epsilon_{ijk}$$

```
> # level-1 variance components
> (c(1.000, coef(mlmm$modelStruct$varStruct, unconstrained = FALSE))*mlmm$sigma)^2
      2      3      4      5      6
0.1009578 0.1637858 0.7404776 0.6789130 0.4454560 0.4557227
> # level-2 error variance (teacher-level student engagement factor variance)
> 0.638**2
[1] 0.407044
> # level-3 error variance (school-level student engagement factor variance)
> 0.418**2
[1] 0.174724
```

Summary

- We ran a multilevel measurement model in an MLM framework, allowing us to recover item means and factor variances simultaneously
- The MLM framework is useful for obtaining latent trait values at multiple levels, conducting uniform differential item functioning analyses, investigating whether item-level predictors vary across clusters and explain response patterns, etc.

Multilevel Measurement Models in an SEM Framework

5

5

MLMMs in SEM Framework

Section Learning Objectives

State the six-step process for conducting an MLMM in an SEM framework

Specify a multilevel CFA

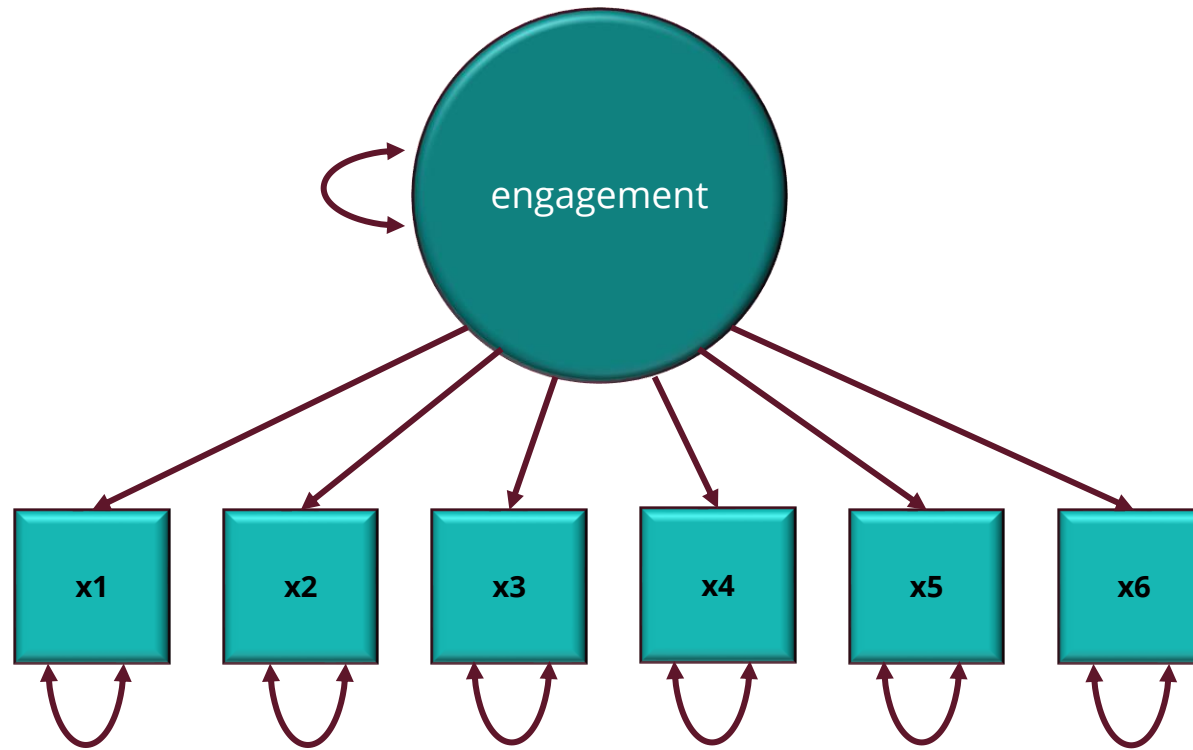
Execute a MLMM in an SEM framework using R package *lavaan* and interpret the output

Compare a single-level CFA to a multilevel CFA

Our Example

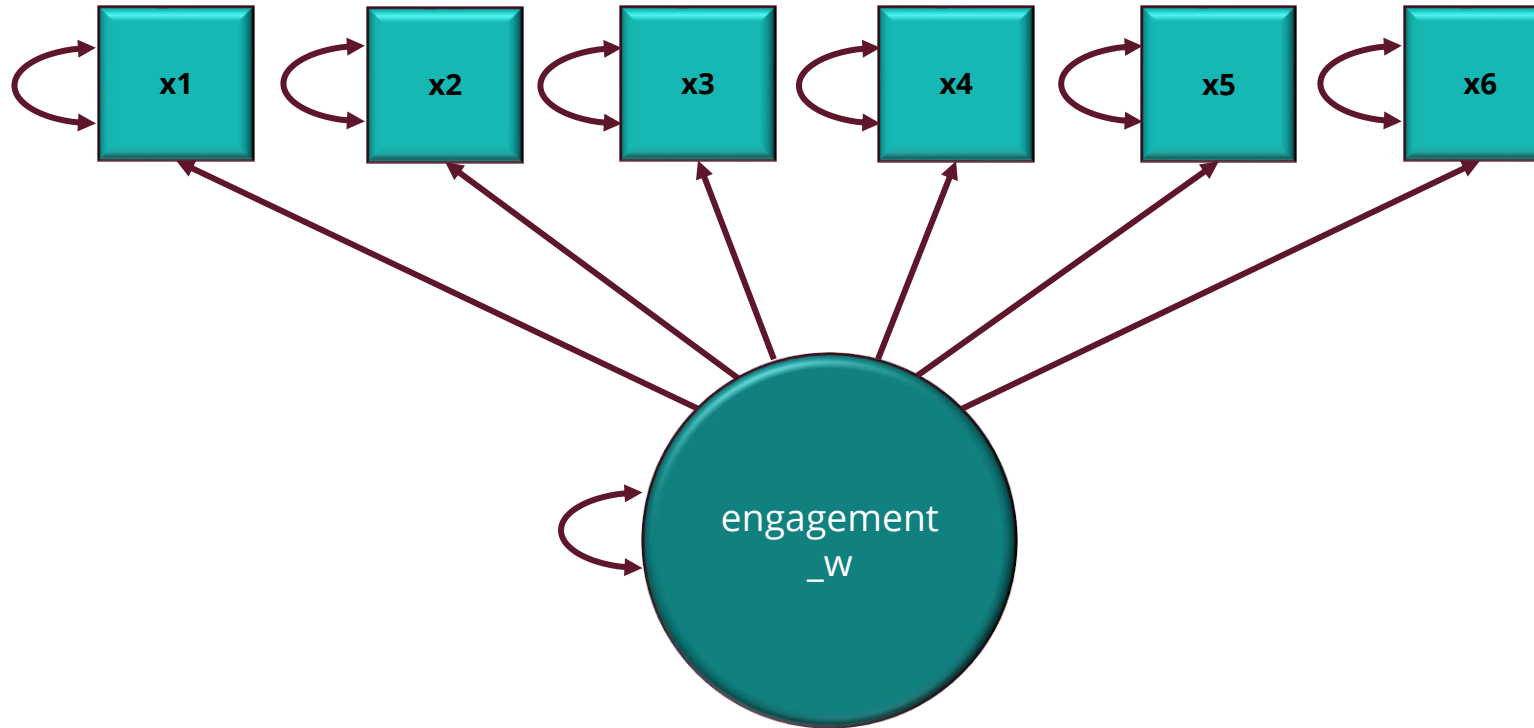
- Data structure: teachers within schools
 - Between 5 and 50 teachers per school
 - 254 schools
- Outcome: perception of student engagement
 - Within: teacher's perception of engagement
 - Between: school-level factor of general engagement
- 6 items assessing student engagement
 - 1 = strongly disagree, 2 = disagree, 3 = somewhat disagree, 4 = somewhat agree, 5 = agree, 6 = strongly agree
- Random subset of data from Huang and Cornell (2015)
 - <https://github.com/flh3/pubdata/blob/main/MLCFA/raw.csv>

MLMM in SEM Framework



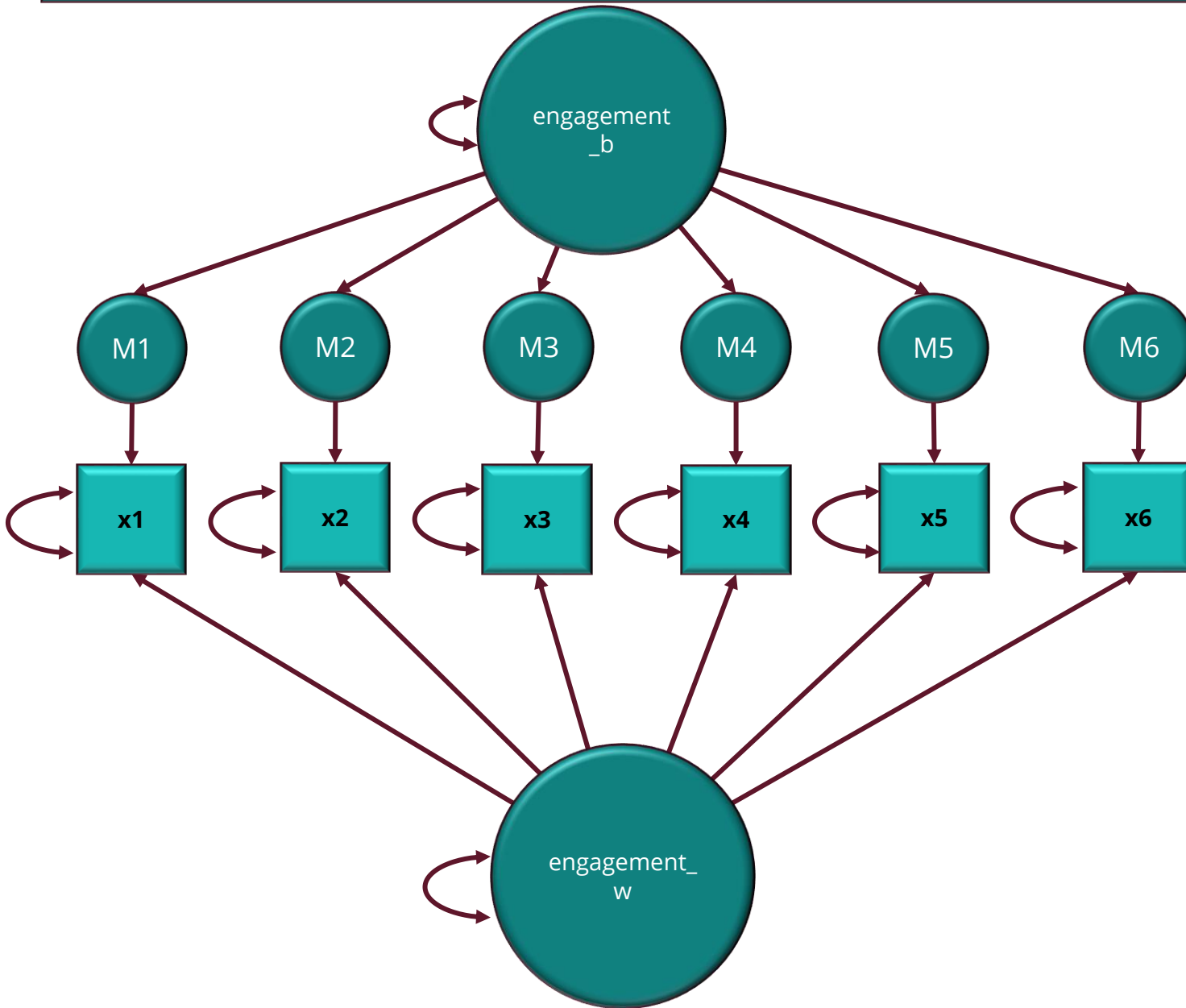
- A regular CFA (pictured) would tell us how our items relate to perceptions of student engagement via the loadings, accounting for measurement variance and factor variance
- This model assumes overall item variances and covariances, doesn't account for potential clustering
- What if our data are clustered?
 - Parameter estimates, standard errors, and model fit might be biased
 - Student engagement might have different meanings at individual and cluster level

MLMM in SEM Framework



- In an MLMM in the SEM framework, the “within” portion of the model looks just like a single-level CFA: items load directly to a within latent variable
- This "within" measurement model captures how the items reflect teacher perceptions of student engagement
- Higher loading = item related to perception of higher student engagement

MLMM in SEM Framework



- Then we add the “between” level, calculating group means that load onto a between latent variable
- The "between" measurement model captures how item means reflect school-level perception of student engagement
- Higher loading = item related to higher levels of school-level perception of engagement
- Within and between covariance matrices estimated simultaneously, partitions mixed total variance into within and between

Steps for Running

Step 1

Prepare Data

Step 2

Check degree of clustering

Step 3

Conventional CFA

Step 4

Within CFA

- Steps 4 and 5 are helpful if you're having convergence issues, generate suggested starting points
- Full data required
 - i.e., not just variance-covariance matrix

Step 5

Between CFA

Step 6

Multilevel CFA

Step 1

Prepare Data

```
library(RCurl) # for extracting data from web
library(lavaan) # for SEM
library(dplyr) # renaming variables
library(lme4) # for ICCs
library(performance) # for ICCs
```

```
data <- read.csv(text = getURL("https://raw.githubusercontent.com/flh3/pubdata/main/MLCFA/raw.csv\"")) %>%
  rename(schoolID = sid)
```


Step 2

Check degree of clustering

- We check the extent of clustering to ascertain whether we need to go through this multi-step process
- If the data aren't very clustered, it might be unnecessary to take the extra steps accounting for clustering
- The intraclass correlation coefficient (ICC) from the multilevel modelling framework is useful here

$$ICC = \frac{\phi_0^2}{\phi_0^2 + \tau_0^2}$$

- With our items as the outcome, the ICC reflects how much variance in item responses is between cluster
- Ranges from 0 to 1, with higher numbers indicating that a larger proportion of item response variance is between clusters
- With $ICC < .05$, multilevel approach may not be needed

Step 2

Check degree of clustering

```
icc_x1 <- lmer(x1 ~ 1 + (1|schoolID), data = data, REML = F)  
performance::icc(icc_x1)
```

```
icc_x2 <- lmer(x2 ~ 1 + (1|schoolID), data = data, REML = F)  
performance::icc(icc_x2)
```

```
icc_x3 <- lmer(x3 ~ 1 + (1|schoolID), data = data, REML = F)  
performance::icc(icc_x3)
```

```
icc_x4 <- lmer(x4 ~ 1 + (1|schoolID), data = data, REML = F)  
performance::icc(icc_x4)
```

```
icc_x5 <- lmer(x5 ~ 1 + (1|schoolID), data = data, REML = F)  
performance::icc(icc_x5)
```

```
icc_x6 <- lmer(x6 ~ 1 + (1|schoolID), data = data, REML = F)  
performance::icc(icc_x6)
```

Output

- x1: 0.203
- x2: 0.258
- x3: 0.143
- x4: 0.109
- x5: 0.240
- x6: 0.127

Step 3

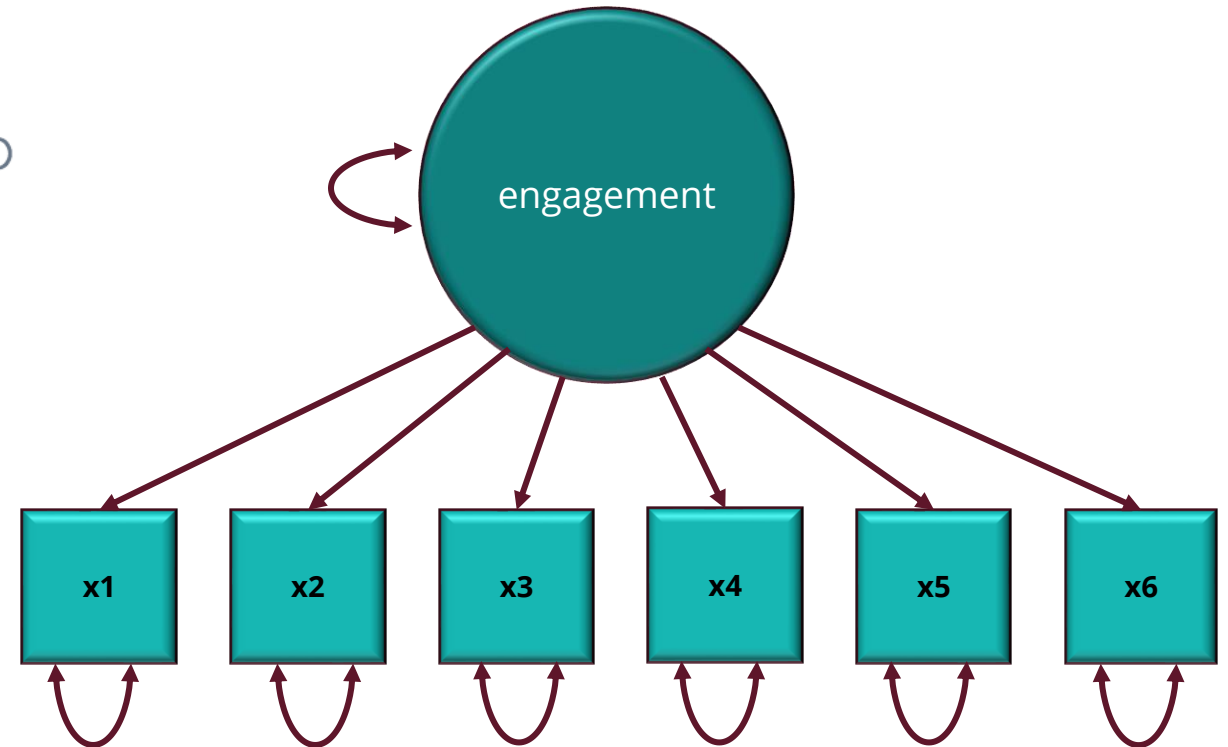
Conventional CFA

- What is it: factor analysis of the total covariance matrix, which results in biased model fit and loading estimates because the total covariance matrix is an uninterpretable blend of within and between effects
- If the results are biased, why conduct this step?
- If your ICC is greater than .05, you don't need to conduct the conventional CFA, this model is more to illustrate the effect of accounting for the clustered structure
- If your ICC is less than .05, you might run a conventional CFA as normal in lieu of using the multilevel approach
 - With insufficient clustering, there is not a lot of mean level variance across clusters, so estimation might be difficult/impossible

Step 3

Conventional CFA

```
regular <- '  
  
  engagement =~ NA*x1 + x2 + x3 + x4 + x5 + x6  
  engagement ~~ 1*engagement  
  
'  
regular_fit <- cfa(model = regular, data = data)  
summary(regular_fit, fit.measures = TRUE, standardized = TRUE)
```



Step 3

Conventional CFA

Model Test User Model:

Test statistic	2226.990
Degrees of freedom	9
P-value (Chi-square)	0.000

Model Test Baseline Model:

Test statistic	12323.196
Degrees of freedom	15
P-value	0.000

User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.820
Tucker-Lewis Index (TLI)	0.700

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-29138.673
Loglikelihood unrestricted model (H1)	-28025.178
Akaike (AIC)	58301.346
Bayesian (BIC)	58376.552
Sample-size adjusted Bayesian (BIC)	58338.421

Root Mean Square Error of Approximation:

RMSEA	0.252
90 Percent confidence interval - lower	0.243
90 Percent confidence interval - upper	0.260
P-value RMSEA \leq 0.05	0.000

Standardized Root Mean Square Residual:

SRMR	0.079
------	-------

Step 3

Conventional CFA

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement =~						
x1	0.736	0.012	62.120	0.000	0.736	0.840
x2	0.840	0.013	63.372	0.000	0.840	0.851
x3	0.700	0.018	39.239	0.000	0.700	0.599
x4	0.550	0.016	33.425	0.000	0.550	0.524
x5	0.804	0.016	50.571	0.000	0.804	0.728
x6	0.821	0.016	51.505	0.000	0.821	0.738

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement	1.000				1.000	1.000
.x1	0.226	0.007	30.699	0.000	0.226	0.294
.x2	0.268	0.009	29.443	0.000	0.268	0.275
.x3	0.876	0.021	41.111	0.000	0.876	0.641
.x4	0.799	0.019	42.095	0.000	0.799	0.725
.x5	0.573	0.015	37.956	0.000	0.573	0.469
.x6	0.564	0.015	37.582	0.000	0.564	0.455

S_{PW} and S_B

- It is possible to run level-specific models
- To do so, you need to manually partition the total variance-covariance matrix into within and between matrices
- S_{PW} (pooled within) is an unbiased estimate of population within covariance matrix
- S_B is a biased estimator of the population between covariance matrix, but can be adjusted with information about average cluster size and S_{PW}

S_{PW} and S_B : Equations

- We'll be using the function written by Huang (2017) to get our partitioned covariance matrices

https://github.com/flh3/mcfa/blob/main/02_syntax/mcfa2.R

- Equations for your reference

$$S_{PW} = (n - G)^{-1} \sum_{g=1}^G \sum_{i=1}^{n_g} (y_{ig} - \bar{y}_g)(y_{ig} - \bar{y}_g)'$$

$$S_B = (G - 1)^{-1} \sum_{g=1}^G n_g (\mathbf{y}_g - \bar{\mathbf{y}})(\mathbf{y}_g - \bar{\mathbf{y}})'$$

$$c. = [n^2 - \sum_{g=1}^G n_g^2][n(G - 1)]^{-1}$$

- $\Sigma_B = \frac{S_B - S_{PW}}{c}$

- n is total sample size
- G is number of groups
- y_{ig} is score of observation i nested in group g
- \bar{y}_g is cluster mean in group g
- \bar{y} is overall grand mean
- c is average cluster size

Huang (2017)
Muthén (1994)

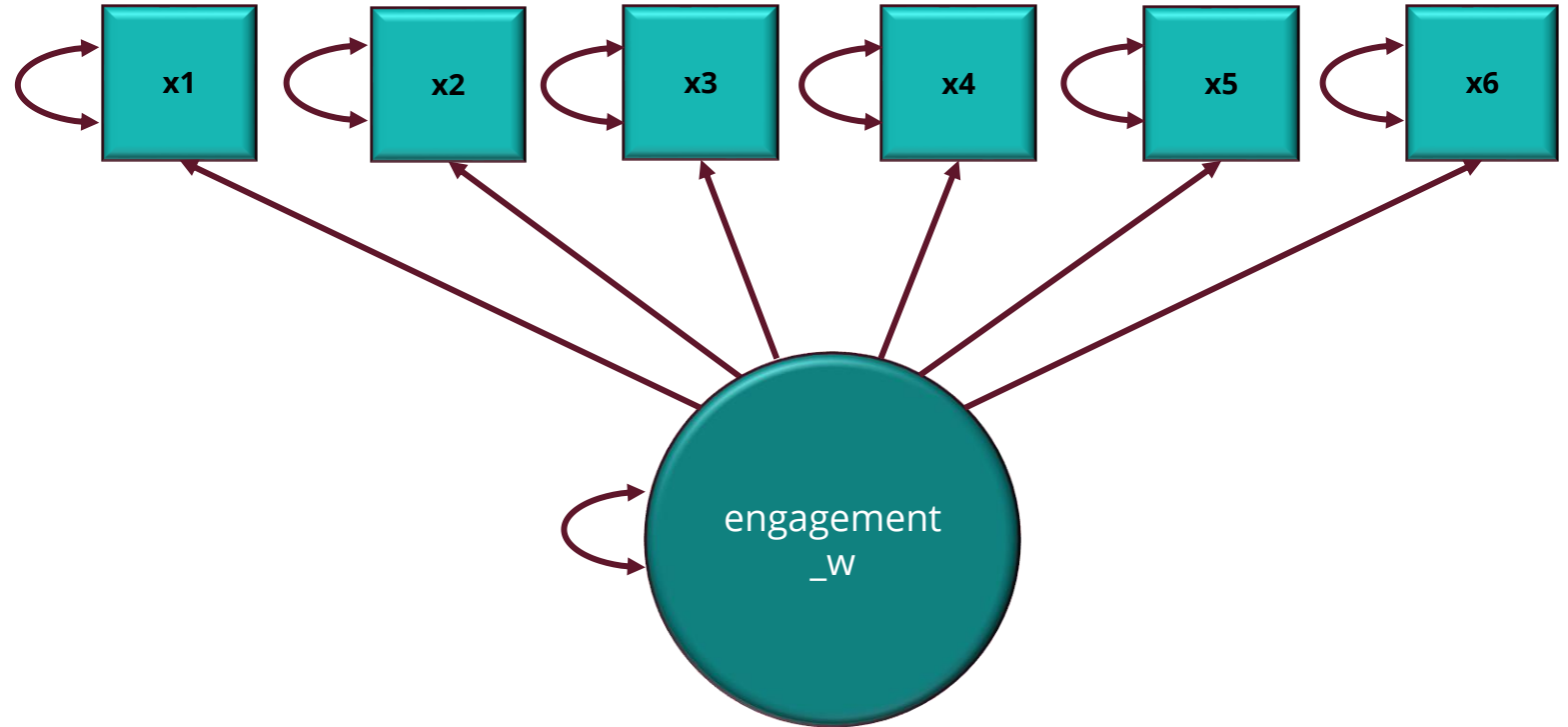
Step 4

Within CFA

- What is it: Factor analysis of within-level using pooled within covariance matrix, S_{PW}
- Why do it: Evaluates the appropriateness of the within structure
- Yields fit information and factor loadings at the within-only level
- If fit is poor, may want to re-evaluate the within model (and document that you did so)
- If you want an unbiased within-only model (i.e., you want to account for clustering but aren't interested in the between model), you can use the output from this model
- Can be used to inform starting values for the MLMM if there are convergence issues (more on that later)

Step 4

Within CFA



```
within <- '  
  
  engagement_w =~ NA*x1 + x2 + x3 + x4 + x5 + x6  
  engagement_w ~~ 1*engagement_w  
  
'  
within_fit <- cfa(within, sample.cov = x$pw.cov, sample.nobs = x$n)  
summary(within_fit, fit.measures = TRUE, standardized = TRUE)
```

Step 4

Within CFA

Loglikelihood and Information Criteria:

Model Test User Model:

Test statistic	2109.180
Degrees of freedom	9
P-value (Chi-square)	0.000

Loglikelihood user model (H0)	-27924.172
Loglikelihood unrestricted model (H1)	-26869.582

Akaike (AIC)	55872.344
Bayesian (BIC)	55947.550
Sample-size adjusted Bayesian (BIC)	55909.420

Model Test Baseline Model:

Test statistic	10000.088
Degrees of freedom	15
P-value	0.000

Root Mean Square Error of Approximation:

RMSEA	0.245
90 Percent confidence interval - lower	0.236
90 Percent confidence interval - upper	0.254
P-value RMSEA ≤ 0.05	0.000

User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.790
Tucker-Lewis Index (TLI)	0.649

Standardized Root Mean Square Residual:

SRMR	0.081
------	-------

Step 4

Within CFA

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_w =~						
x1	0.583	0.011	50.692	0.000	0.583	0.743
x2	0.642	0.013	51.300	0.000	0.642	0.749
x3	0.633	0.017	37.314	0.000	0.633	0.585
x4	0.450	0.016	27.759	0.000	0.450	0.454
x5	0.710	0.014	50.542	0.000	0.710	0.741
x6	0.800	0.015	53.102	0.000	0.800	0.768

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_w	1.000				1.000	1.000
.x1	0.276	0.008	34.740	0.000	0.276	0.448
.x2	0.322	0.009	34.356	0.000	0.322	0.439
.x3	0.769	0.019	40.216	0.000	0.769	0.658
.x4	0.780	0.018	42.190	0.000	0.780	0.794
.x5	0.414	0.012	34.833	0.000	0.414	0.451
.x6	0.444	0.013	33.110	0.000	0.444	0.410

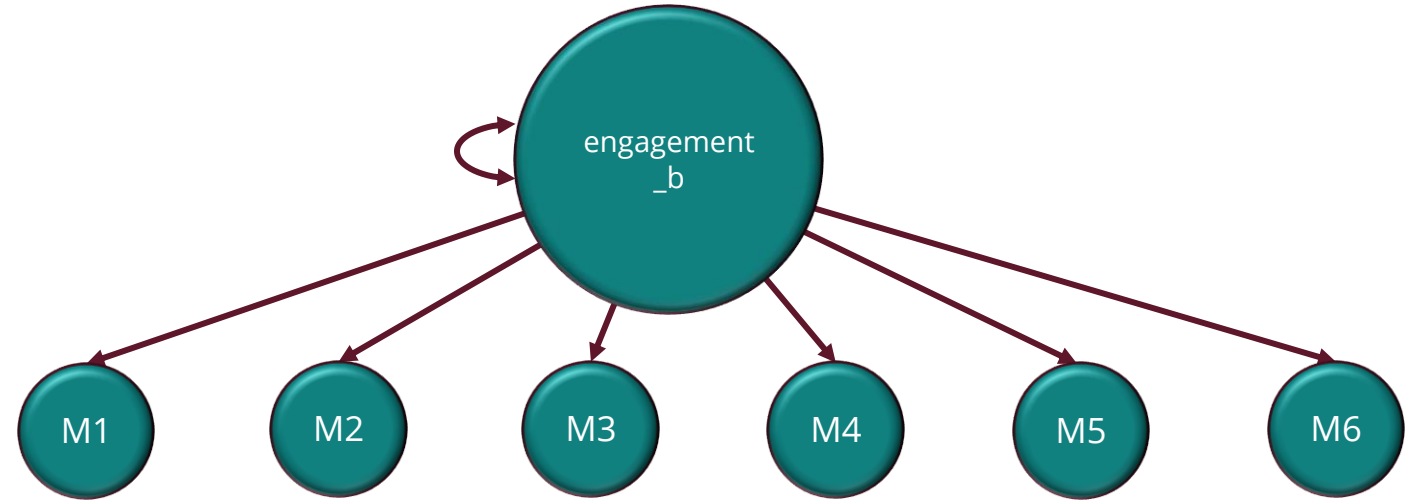
Step 5

Between CFA

- What is it: Factor analysis of between level using between-group covariance matrix, Σ_B
- Why do it: Evaluates appropriateness of between structure
- Yields model fit and loadings at between level
- Might consider changing structure if fit is poor (and, again, documenting that you did so)
 - EFA with between-group covariance matrix
- Like within-only model, can be used to inform starting values for the MLMM if there are convergence issues

Step 5

Between CFA



```
between <- '
```

```
engagement_b =~ NA*x1 + x2 + x3 + x4 + x5 + x6  
engagement_b ~~ 1*engagement_b
```

```
,'  
between_fit <- cfa(between, sample.cov = x$b.cov, sample.nobs = x$G)  
summary(between_fit, fit.measures = T, standardized = T)
```

Step 5

Between CFA

Model Test User Model:

Test statistic	257.623
Degrees of freedom	9
P-value (Chi-square)	0.000

Model Test Baseline Model:

Test statistic	1706.279
Degrees of freedom	15
P-value	0.000

User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.853
Tucker-Lewis Index (TLI)	0.755

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-2450.816
Loglikelihood unrestricted model (H1)	-2322.005
Akaike (AIC)	4925.632
Bayesian (BIC)	4968.080
Sample-size adjusted Bayesian (BIC)	4930.038

Root Mean Square Error of Approximation:

RMSEA	0.330
90 Percent confidence interval - lower	0.296
90 Percent confidence interval - upper	0.365
P-value RMSEA \leq 0.05	0.000

Standardized Root Mean Square Residual:

SRMR	0.071
------	-------

Step 5

Between CFA

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_b =~						
x1	1.643	0.080	20.553	0.000	1.643	0.956
x2	2.027	0.097	20.910	0.000	2.027	0.965
x3	1.564	0.109	14.318	0.000	1.564	0.764
x4	1.260	0.090	14.030	0.000	1.260	0.753
x5	1.901	0.122	15.568	0.000	1.901	0.809
x6	1.573	0.094	16.763	0.000	1.573	0.848

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_b	1.000				1.000	1.000
.x1	0.253	0.037	6.846	0.000	0.253	0.086
.x2	0.303	0.052	5.870	0.000	0.303	0.069
.x3	1.747	0.162	10.760	0.000	1.747	0.417
.x4	1.212	0.112	10.793	0.000	1.212	0.433
.x5	1.913	0.181	10.578	0.000	1.913	0.346
.x6	0.963	0.093	10.322	0.000	0.963	0.280

Step 6

Multilevel CFA

- What is it: simultaneous factor analysis of both within- and between-group covariance matrices
- Why do it: model building all in one place is easier, and allows for ease of building on a structural model
- What will it tell you: loadings and fit information for within and between models simultaneously

Step 6

Multilevel CFA

```
multilevel <- '
```

```
  level: 1
```

```
    engagement_w =~ NA*x1 + x2 + x3 + x4 + x5 + x6
```

```
    engagement_w ~~ 1*engagement_w
```

```
  level: 2
```

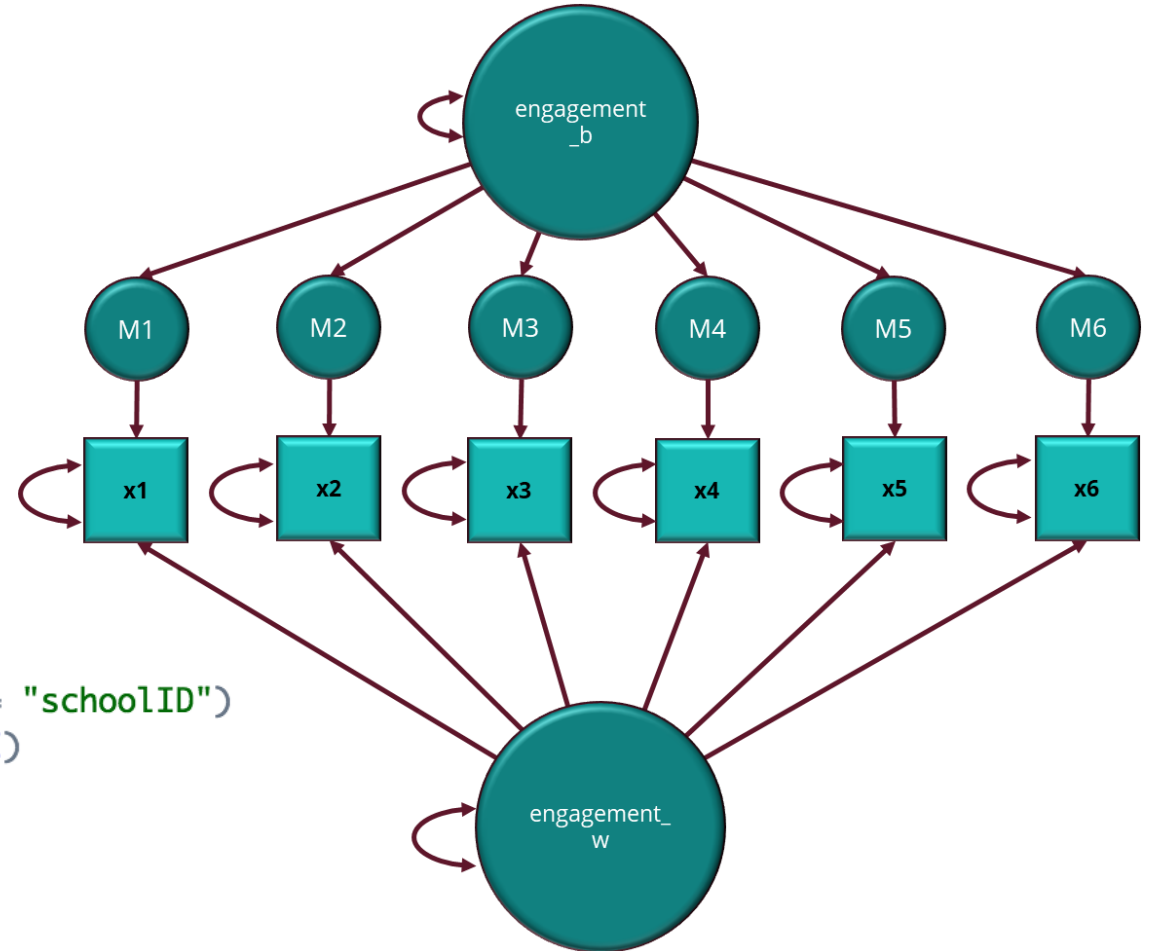
```
    engagement_b =~ NA*x1 + x2 + x3 + x4 + x5 + x6
```

```
    engagement_b ~~ 1*engagement_b
```

```
'
```

```
multilevel_fit <- cfa(model = multilevel, data = data, cluster = "schoolID")
```

```
summary(multilevel_fit, fit.measures = TRUE, standardized = TRUE)
```



Step 6

Multilevel CFA

Model Test User Model:

Test statistic	2207.215
Degrees of freedom	18
P-value (Chi-square)	0.000

Model Test Baseline Model:

Test statistic	10988.837
Degrees of freedom	30
P-value	0.000

User Model versus Baseline Model:

Comparative Fit Index (CFI)	0.800
Tucker-Lewis Index (TLI)	0.667

Loglikelihood and Information Criteria:

Loglikelihood user model (H0)	-28526.495
Loglikelihood unrestricted model (H1)	-27422.888
Akaike (AIC)	57112.991
Bayesian (BIC)	57301.007
Sample-size adjusted Bayesian (BIC)	57205.680

Root Mean Square Error of Approximation:

RMSEA	0.177
90 Percent confidence interval - lower	0.171
90 Percent confidence interval - upper	0.183
P-value RMSEA \leq 0.05	0.000

Standardized Root Mean Square Residual (corr metric):

SRMR (within covariance matrix)	0.082
SRMR (between covariance matrix)	0.083

Step 6

Multilevel CFA

Level 1 [within]:

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_w =~						
x1	0.554	0.015	37.942	0.000	0.554	0.709
x2	0.618	0.016	39.443	0.000	0.618	0.721
x3	0.664	0.019	35.735	0.000	0.664	0.609
x4	0.437	0.017	25.388	0.000	0.437	0.440
x5	0.744	0.016	45.134	0.000	0.744	0.770
x6	0.837	0.018	46.600	0.000	0.837	0.797

Intercepts:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
.x1	0.000				0.000	0.000
.x2	0.000				0.000	0.000
.x3	0.000				0.000	0.000
.x4	0.000				0.000	0.000
.x5	0.000				0.000	0.000
.x6	0.000				0.000	0.000
engagement_w	0.000				0.000	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_w	1.000				1.000	1.000
.x1	0.303	0.012	24.479	0.000	0.303	0.497
.x2	0.353	0.015	24.274	0.000	0.353	0.480
.x3	0.746	0.020	36.689	0.000	0.746	0.629
.x4	0.799	0.020	40.303	0.000	0.799	0.807
.x5	0.381	0.015	24.598	0.000	0.381	0.407
.x6	0.403	0.019	21.618	0.000	0.403	0.365

Step 6

Multilevel CFA

Level 2 [schoolID]:

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_b =~						
x1	0.410	0.023	17.766	0.000	0.410	1.006
x2	0.503	0.028	18.274	0.000	0.503	1.006
x3	0.326	0.031	10.528	0.000	0.326	0.802
x4	0.308	0.025	12.268	0.000	0.308	0.911
x5	0.395	0.034	11.783	0.000	0.395	0.799
x6	0.339	0.027	12.377	0.000	0.339	0.922

Intercepts:

	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
.x1	4.743	0.029	163.740	0.000	4.743	11.636
.x2	4.562	0.035	130.713	0.000	4.562	9.117
.x3	3.518	0.032	110.771	0.000	3.518	8.672
.x4	4.348	0.027	160.013	0.000	4.348	12.877
.x5	4.202	0.035	118.453	0.000	4.202	8.504
.x6	3.917	0.029	133.737	0.000	3.917	10.655
engagement_b	0.000				0.000	0.000

Variances:

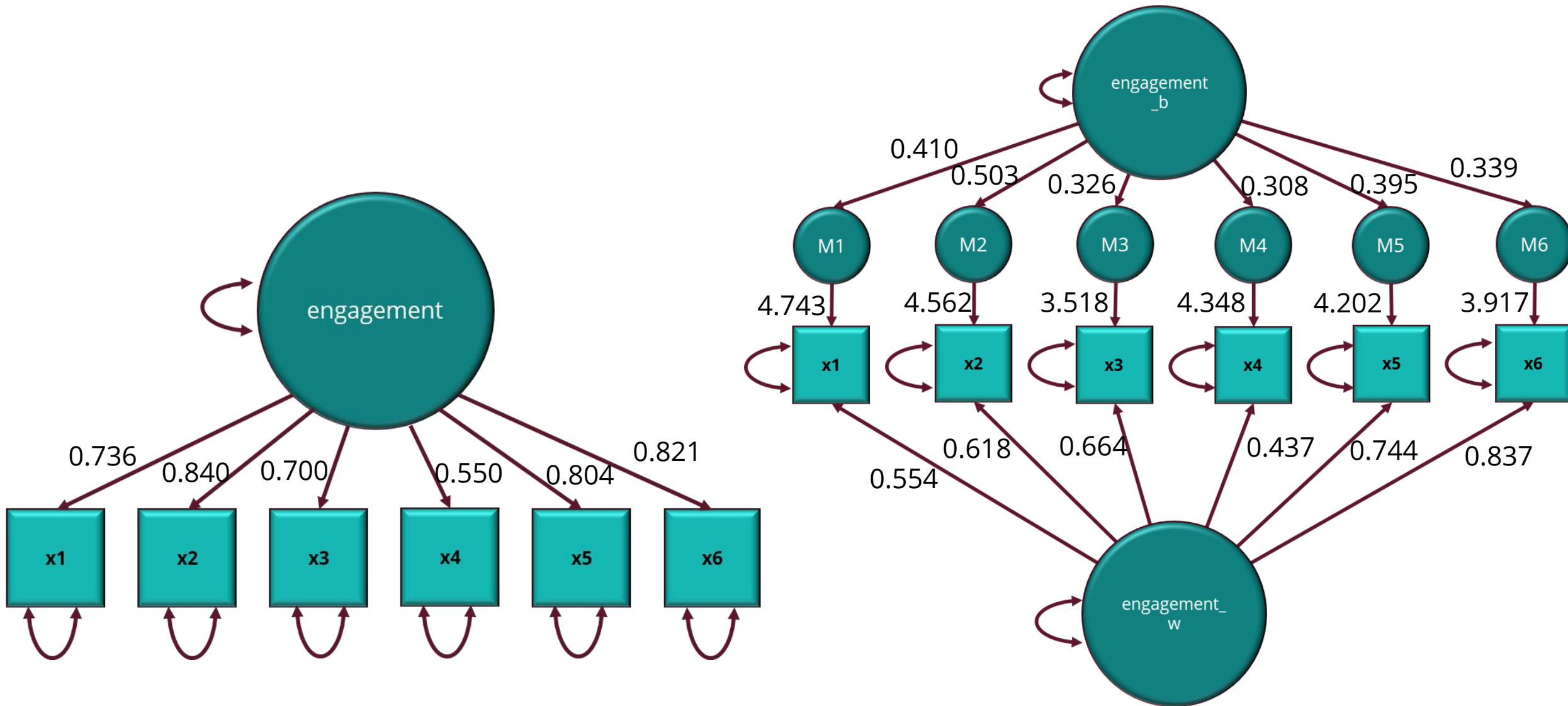
	Estimate	Std.Err	z-value	P(> z)	Std.lv	Std.all
engagement_b	1.000				1.000	1.000
.x1	-0.002	0.002	-1.066	0.286	-0.002	-0.012
.x2	-0.003	0.003	-1.033	0.302	-0.003	-0.012
.x3	0.059	0.011	5.459	0.000	0.059	0.356
.x4	0.019	0.008	2.540	0.011	0.019	0.170
.x5	0.088	0.011	7.771	0.000	0.088	0.361
.x6	0.020	0.006	3.497	0.000	0.020	0.150

Variances can't be negative, these are likely near-zero

Comparing CFA Results

Models	χ^2	df	CFI	RMSEA	SRMR
Regular	2226.990	9	0.820	0.700	0.079
Within	2109.180	9	0.790	0.245	0.081
Between	257.623	9	0.853	0.330	0.071
Multilevel	2207.215	18	0.800	0.177	W = 0.082 B = 0.083

Comparing Regular and Multilevel CFA Results



Setting Starting Values

- As mentioned, you might want to run the individual-level models if your initial MLMM doesn't converge to set starting values
- To set starting values, you use the special *start()* function in your model specification

```
multilevel <- '  
  
  level: 1  
    engagement_w =~ NA*x1 + start(0.642)*x2 + start(0.633)*x3 +  
                      start(0.450)*x4 + start(0.710)*x5 + start(0.800)*x6  
    engagement_w ~~ 1*engagement_w  
  
  level: 2  
    engagement_b =~ NA*x1 + start(2.027)*x2 + start(1.564)*x3 +  
                      start(1.260)*x4 + start(1.901)*x5 + start(1.573)*x6  
    engagement_b ~~ 1*engagement_b  
  
'
```

- The values input to *start()* are those from the individual-level models

Summary

- Multilevel measurement modelling in an SEM framework is more flexible than in an MLM framework
- A multi-step process is recommended for understanding your modelling: check clustering, regular CFA, within-only, between-only, multilevel
- To conduct a multilevel CFA, the total covariance matrix is partitioned into within and between covariance matrices, either manually or automatically with software
- There is a lot of room for decision-making, so documenting your process is important

You have reached the end of this section...

Shaw, M. & Flake, J. K. (2023). Multilevel Measurement Models [Digital ITEMS Module 34]. *Educational Measurement: Issues and Practice*, 34(4), 82.